

1 **Evaluation of n-tree distance sampling for inventory of headwater riparian**
2 **forests of western Oregon.**

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4 **Abstract**

5 n-Tree distance sampling (NTDS), also known as k-tree sampling and point-to-tree sampling, has
6 been promoted as a practical method for forest inventory. This simulation study evaluated the
7 performance of three NTDS estimators, as compared with fixed plot sampling and horizontal
8 point sampling, for estimating density and basal area in headwater riparian forests of western
9 Oregon. Bias of at least one NTDS estimator was low for both density and basal area when
10 at least six trees were captured at each sample point, but performance of NTDS for density
11 estimation was poor on stem maps exhibiting a clustered pattern. We close with some comments
12 regarding the statistical efficiency of NTDS for riparian area inventory in similar forest
13 conditions.

14 **Keywords**

15 Forest sampling, density-adapted sampling, *k*-tree sampling, Pacific Northwest
16

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17 **Introduction**

18 Located on the fringes of the drainage network, headwater streams are intimately connected with
19 downstream reaches, serving as a source of sediment, woody debris, organic matter and nutrients
20 (MacDonald and Coe 2007). Nonetheless, forests adjoining non-fish-bearing (Type N) streams in
21 western Oregon receive no legal protection from timber harvest (Adams 2007). Therefore, forest
22 managers have the opportunity—and responsibility—to actively manage headwater riparian
23 systems for a variety of wildlife habitat, watershed protection, and fiber production objectives.

24 Accurate and efficient estimation of stem density and basal area of trees on an area (e.g.
25 per-acre) basis can be crucial to the success of active restoration or management programs in
26 forests adjacent to headwater streams. Nonetheless, inventory in riparian forests can be more
27 difficult than in their upland counterparts. Stand structure and composition can be highly variable
28 (Pabst and Spies 1999) as a result of hydrologic disturbance and other fine-scale processes.
29 Particularly in naturally regenerated areas, headwater streams can contain alternating patches of
30 conifer and hardwood trees of varying sizes. The development of new inventory methods,
31 specifically designed to mitigate these challenges, would be a welcome addition to the forest
32 sampling toolbox.

33 Most forest inventories in the Pacific Northwest are conducted using fixed plot and/or
34 horizontal point sampling designs. Circular fixed plot sampling (FPS) is one of the oldest
35 methods of forest sampling, and is still commonly used throughout much of the world.
36 Horizontal point sampling (HPS), commonly known as variable plot sampling (VPS) in the
37 Pacific Northwest, was developed by W. Bitterlich in 1948 and introduced to North American
38 foresters by Grosenbaugh (1952). Under FPS, density can be estimated with a simple count of
39 “in” trees, but basal area estimation requires diameter measurements on at least some captured

40 trees. The exact opposite is true under HPS. These practical concerns, combined with the
41 efficiency gained when selection probability is made proportional to the attribute of interest
42 (Grosenbaugh 1967), tend to make FPS more statistically efficient for density estimation and
43 HPS more efficient for basal area estimation. N-tree-distance sampling (NTDS), also called k -
44 tree sampling, density-adapted sampling, point-to-tree sampling or simply “distance sampling”,
45 was promoted as an “all-encompassing forest inventory method” by Jonsson et al. (1992). In this
46 method, the n trees nearest the sample point are selected (n being a pre-determined number that
47 remains constant throughout the sampling effort). Since the same number of trees is captured at
48 all sample points, empty plots and plots with too many trees can be avoided (Kleinn and Vilčko
49 2006a), leading to a potential increase in productivity. In addition, the distance to the center of
50 the n tree, acquired as a byproduct of this system, may provide some information on the spatial
51 distribution of trees within a forest (Lessard et al. 1994).

52 One important drawback of NTDS is that, unlike FPS and HPS, selection probabilities of
53 individual trees cannot be known unless distances and azimuths to many additional trees are
54 acquired (Kleinn and Vilčko 2006b). Therefore, design-unbiased estimation for this method is
55 currently not operationally feasible. A plot area for all n trees can be computed as a circle with a
56 radius that is equal to the horizontal distance to the center of the n tree. The factor used to expand
57 per-plot estimates to a per-acre basis is then $EF = 43,560/A_p$, where A_p is the plot area in ft^2 .
58 Because the plot size at each sample point is computed as the smallest that could contain n trees,
59 this uncorrected estimator will systematically overestimate the value of any attribute on a per-
60 acre basis (Kleinn and Vilčko 2006a).

61 Despite the lack of a practical, design-unbiased estimator for NTDS, previous authors
62 have promoted it as an attractive sampling method on the basis of an ease in field application

63 (Jonsson et al. 1992; Kleinn and Vilčko 2006b; Nothdurft 2010). We hypothesized that the
64 ability to control the number of trees captured at each sample point may make NTDS an
65 attractive option for inventorying highly variable riparian stands. HPS has historically been
66 preferred by many forest inventory professionals in the Pacific Northwest because, when a prism
67 or Relascope is employed, horizontal distance measurements are unnecessary except to check the
68 in/out status of borderline trees – a great asset when working in steep and brushy terrain. The
69 advent of portable rangefinders with ever more sophisticated brush-filtering capacity has
70 increased the viability of alternative sampling systems that depend more directly on distance
71 measurements. Now that NTDS is more technologically feasible, we hoped to explore the
72 relative advantages and disadvantages of the method in a practical sampling application.

73 The objective of this study was to examine the performance of selected NTDS estimators
74 for estimation of density and basal area of headwater riparian forests in western Oregon. We
75 evaluated the NTDS estimators against each other, but also against FPS and HPS for both density
76 and basal area estimation. In the last 40 years, much effort has gone toward the development of
77 estimators for NTDS that minimize bias under a range of forest conditions (see Magnussen 2008
78 for a good overview). However, many of these estimators are difficult to comprehend and
79 implement without advanced statistical training, and are therefore inaccessible to the majority of
80 forest inventory professionals. Three estimators were chosen on the basis of their simplicity in
81 understanding and application, as well as their track record in previous studies. These will be
82 termed the Moore (Moore 1954), Prodan (as described in Lynch and Rusydi 1999) and Kleinn-
83 Vilčko (Kleinn and Vilčko 2006a) estimators. Details on the computation and theoretical
84 background of these estimators can be found in the appendix.

85 **Methods**

86 Data were collected at eight different riparian sites as part of the Bureau of Land
87 Management Density Management and Riparian Buffer Study, an interdisciplinary study on the
88 effect of management activities on wildlife habitat and other ecosystem attributes of riparian and
89 upland systems (Cissel et al. 2006). At each site, a 1.28-ac square plot was established so as to
90 have an approximately equal area on both sides of the stream, within which the species, dbh and
91 coordinate position of every tree was recorded (see Marquardt et al. 2010 for details regarding
92 inventory procedures). Sites were dominated by Douglas-fir (*Pseudotsuga menziesii* var.
93 *menziesii* [Mirb.] Franco) or western hemlock (*Tsuga heterophylla* [Raf.] Sarg.), with western
94 redcedar (*Thuja plicata* Donn ex D. Don) a minor component of some sites. Hardwood species
95 such as red alder (*Alnus rubra* Bong.) and bigleaf maple (*Acer macrophyllum* Pursh) were
96 present at most sites. Composition of each site, by density and basal area, is given in Table 1.

97 Because the performance of NTDS estimation has been found in previous work (e.g.
98 Lessard et al. 1994; Kleinn and Vilčko 2006a) to be highly dependent on the spatial distribution
99 of the trees on the tract of interest, the spatial distribution of each site was quantified, by species
100 and for the site as a whole, using the Clark-Evans (CE) index (Clark and Evans 1954). The CE
101 index takes on values of: 0 if the population is extremely aggregated (i.e. clustered); 1 if the
102 population is distributed completely at random; and 2.14 if the population is perfectly uniform.
103 Computation of the CE index was done using the program SIAFOR (Kint et al. 2004). Results
104 are shown in Table 2.

105 The estimators were compared using a Monte Carlo sampling algorithm written in the
106 Microsoft Visual Basic for Applications (VBA) programming language (v. 6.5, Microsoft
107 Corporation, Redmond, WA). Following Kleinn and Vilčko (2006a), in order to provide a
108 common basis for evaluation, the estimators were compared across a range of n (the desired

109 number of trees per sample point) from 2 to 10. While a certain number of trees captured does
110 not necessarily represent an equal amount of effort across sampling methods (e.g. for a given
111 number of trees, HPS requires less measurement time for basal area estimation than NTDS), we
112 believe that the relative performance for different values of n will allow rough comparison of
113 different sampling systems and more specific comparison between NTDS estimators.

114 The value of n was fixed for each simulation run. The FPS radius and HPS basal area
115 factor were computed so that n trees per sample point would be captured on average. For FPS,
116 the plot area A_{FPS} , in ac, was set as $A_{FPS} = n/TPA$, where TPA is the density (in trees/ac) of the
117 stem map. For HPS, the basal area factor, in ft^2/ac , was set as $BAF = BA/n$, where BA is the basal
118 area (ft^2/ac) of the stem map. Toroidal wrapping, which gives all trees the appropriate long-run
119 probability of selection, was used to avoid underselection problems associated with edge effects.

120 At each iteration of the simulation, one sample point was randomly located using pseudo-
121 random numbers generated by VBA. Estimates of density and basal area were computed for each
122 estimation method simulated (FPS, HPS and each NTDS estimator). This process was repeated
123 10,000 times, with a different seed set for each iteration in order to avoid cyclical number
124 generation patterns. A sample size of 10,000 seemed to be adequate for characterizing the
125 statistical performance of each estimation method, as evidenced by the low realized values of the
126 maximum recorded relative bias of FPS for density estimation and HPS for basal area estimation
127 (neither was greater than 3%).

128 The performance of each estimation method was evaluated using relative bias and
129 relative root mean square error (RRMSE). Relative values were preferred because they allow an
130 equal basis of comparison between attributes and sites. Relative bias was computed as:

131
$$RB = \frac{(\bar{Y} - Y) * 100}{Y}$$

132 where Y is the true value of density or basal area, $\bar{Y} = \sum_{i=1}^{10,000} \hat{Y}_i / 10,000$ is the mean estimate and
133 \hat{Y}_i is the estimate produced at iteration i .

134 RRMSE was computed as:

$$135 \quad RRMSE = \sqrt{\frac{\sum_{i=1}^{10,000} [(\hat{Y}_i - \bar{Y})^2]}{(10000 - 1) * N} + (\bar{Y} - Y)^2 \left(\frac{100}{Y}\right)}$$

136 where N is the sample size.

137 **Results**

138 *Relative bias*

139 Because they are design-unbiased for estimation of density and basal area, bias results for
140 FPS and HPS will not be presented. Among NTDS estimators, the Moore estimator tended to
141 underestimate density (Figure 1) and basal area (Figure 2) for small values of n , while the Prodan
142 and Kleinn-Vilčko estimators tended to give upwardly biased estimates. For a given value of n ,
143 the Moore estimator clearly had the lowest absolute relative bias for density estimation at seven
144 sites, particularly for $n \geq 4$. For basal area estimation, the Moore estimator clearly had the lowest
145 absolute relative bias at six sites, with the identity of the lowest-bias estimator unclear at two
146 sites. For density and basal area estimation, the Prodan estimator had the highest bias for small
147 values of n , but appeared to converge with the Kleinn-Vilčko estimator for $n > 5$.

148 *RRMSE*

149 RRMSE was calculated across a range of N for a moderate value of $n=6$. For estimation
150 of density, FPS had the lowest RRMSE across all values of N (Figure 3). HPS and the Moore
151 estimator had roughly equal performance across most sites. A gap between the estimation

152 methods previously mentioned and the Prodan and Kleinn-Vilčko estimators was evident across
153 all sites, with the latter giving notably poor performance.

154 For estimation of basal area, HPS had the lowest RRMSE across all values of N (Figure
155 4), although there was not much difference in performance between HPS, FPS and the Moore
156 estimator. There was a gap in performance between these estimation methods and the Prodan and
157 Kleinn-Vilčko estimators at most (though not all) sites.

158 **Discussion**

159 One limitation in this study is inherent in the use of toroidal wrapping as an edge-effect
160 correction measure. As toroidal wrapping causes sample points located at the edges of the stem
161 map to “wrap around” to the opposite side of the stem map, the simulated RRMSE values can be
162 different from those that would be obtained if (in the most ideal case) the study were conducted
163 on much larger stem maps with a buffer zone surrounding a smaller 1.28-ac area wherein sample
164 points were allowed to fall. The difference can be exacerbated by larger values of n (translating
165 to larger inclusion areas across all estimation methods), which cause the toroidal wrapping to be
166 employed more frequently. Table 3 indicates that the plot size for $n=6$ under FPS can be as large
167 as 0.045 ac, while the largest inclusion area under HPS (corresponding to the largest-diameter
168 tree on the stem map) for $n=6$ can be as large as 0.244 ac. It is unlikely that the use of toroidal
169 wrapping would result in a distorted comparison of RRMSE values among the three NTDS
170 estimators examined, as all will have similar inclusion areas. However, the simulated RRMSE
171 values should be taken with a slight grain of salt when the NTDS estimators are compared with
172 FPS and (especially) HPS.

173 The three NTDS estimators were chosen for this study on the basis of their simplicity and
174 ease of application. The development of new NTDS estimators is an area of current research in

175 Canada and Europe (Nothdurft et al. 2010; Magnussen et al. 2008), and future estimators may
176 have different statistical properties than those evaluated in this study. Therefore, the comments
177 made about the properties of the NTDS estimators evaluated here should not be extended to
178 other NTDS estimators that were not evaluated.

179 The appeal of computer-based simulation studies, such as the one reported here, is that
180 they allow researchers to compare the performance of different forest sampling methods without
181 the expense of fieldwork. However, they are unable to directly compare the statistical efficiency
182 (that is, the precision gained for a given cost investment; Iles 2003, p28) of different sampling
183 methods because the per-sample-point costs of the various methods being compared is not
184 precisely known. The best way to compare the relative statistical efficiency of different sampling
185 methods in a specific setting is through a timed field trial employing experienced cruising staff,
186 preferably in a tract in which the true density and basal area are known. As such a study has yet
187 to be performed in this forest type, we must augment data with conjecture and experience in
188 order to offer suggestions as to how the NTDS estimators examined might compare with HPS
189 and FPS for inventory work in headwater riparian forests of western Oregon.

190 The Moore estimator emerged as the best candidate among the NTDS estimators
191 examined. The Moore estimator had the lowest bias for estimation of density and basal area on
192 most stem maps. Similarly, the Moore estimator had the lowest RRMSE values on most stem
193 maps when larger sample sizes were considered, particularly for estimation of density. The
194 Prodan estimator performed poorly on the stem maps examined. In evaluating the same NTDS
195 estimators, Kleinn and Vilčko (2006a) found that the Prodan estimator had the highest bias on all
196 stem maps but those with a uniform spatial distribution. Lynch and Rusydi (1999) found that the
197 Prodan estimator had negligible bias in uniformly-spaced teak plantations, where the Moore

198 estimator tended to underestimate volume and density. The poor performance of the Prodan
199 estimator in this study may be due to the non-uniform spatial distribution of trees in the sites
200 examined. The performance of the Kleinn-Vilčko estimator was intermediate between that of the
201 Moore and Prodan estimators. Kleinn and Vilčko (2006a) found that the Kleinn-Vilčko estimator
202 had higher bias than the Moore estimator, which they refer to as the Eberhardt estimator, for
203 estimation of density and basal area on most stem maps. However, in contrast to this study, they
204 did not find substantial differences in root squared error (similar to the RRMSE statistic)
205 between the two estimators.

206 As a measure of statistical performance, RRMSE incorporates both the standard error of
207 the sample mean (which decreases as sample size increases) and the bias of the sample mean
208 (which is not affected by sample size). As predicted by theory, FPS and HPS had the the lowest
209 RRMSE values for estimation of density and basal area, respectively. For estimation of basal
210 area, RRMSE values for the Moore estimator and FPS appeared to converge with those of HPS
211 with increasing sample size. For estimation of density, RRMSE values for the Moore estimator
212 and HPS appeared to likewise converge with those of FPS on most stem maps. The increasing
213 competitiveness of FPS and HPS for basal area and density estimation, respectively, with larger
214 sample size is reflective of the design-unbiasedness of these sampling methods. Similarly, the
215 competitiveness exhibited by the Moore estimator was due in part to low absolute bias on most
216 stem maps for reasonable values of n (≥ 4).

217 In timed field trials, Lessard et al. (1994) and Lynch and Rusydi (1999) found NTDS to
218 be cost-competitive with FPS for density estimation in two very different landscapes (northern
219 hardwoods and red pine forests of northern Michigan and Indonesian teak plantations,
220 respectively). Since density estimation for the NTDS estimators examined only requires

221 measurement of the distance to the n tree (and, if implementing the Kleinn-Vilčko estimator, the
222 distance to the $n+1$ tree as well), NTDS may be most promising as a sampling method for
223 density estimation.

224 However, HPS and the Moore estimator gave notably poor performance for density
225 estimation at TH75, the only site with a large hardwood component. At this site bigleaf maple
226 comprised 25% of total stems, but only 8% of total basal area. Red alder was also present,
227 comprising 2% of all stems and 2% of total basal area.

228 The gap in performance appears to stem from the clustered nature of the hardwood
229 species. Bigleaf maple had a CE index of 0.58 and red alder had a CE index of 0.24, indicating a
230 strong tendency towards a clustered spatial distribution for both species. When the distance to
231 the n tree is extremely small (as might happen when a sample point is located inside a hardwood
232 clump), a large overestimate of stem density can be produced. Similarly, the clustered
233 distribution of the hardwood trees, in combination with their small size relative to the population
234 as a whole (the quadratic mean diameter of hardwood trees was 8 in., while the overall quadratic
235 mean diameter was 14 in.) may have contributed to the poor performance of HPS at this site.

236 In timed trials in mixed pine-hardwood forests of southern Maine and New Hampshire,
237 Kenning et al. (2005) found that the Moore estimator (although not referred to by that name)
238 consistently underestimated snag density for very low values of n (one, two and three-tree
239 sampling), a result that is consistent with our findings regarding the underestimation bias of the
240 Moore estimator for very low values of n . They found that the Moore estimator did not offer
241 productivity gains sufficient to compensate for the loss of design-unbiasedness, although a
242 “distance-limited” modification of the Moore estimator showed some promise. They mentioned
243 that snags in the compartments they sampled exhibited a clustered spatial distribution, and

244 indicated that this was at least partly due to “the abundance of dead sprouts of *Acer rubrum*.”
245 Similarly, Lessard et al. (1994) found that the Moore estimator performed poorly in the
246 “clumped, mixed hardwood stands” of northern Michigan.

247 For basal area estimation, RRMSE of the Moore estimator converged with that of FPS
248 and HPS with larger values of N at most sites. However, it seems difficult to imagine that any
249 NTDS estimator could be more statistically efficient than HPS for basal area estimation, since
250 trees must be measured for diameter under NTDS but only counted in HPS. Nonetheless, Lessard
251 et al. (1994) found that, for a fixed sampling time, NTDS using the Moore estimator sometimes
252 gave a lower sampling error than HPS. In another trial, Lynch and Rusydi (1999) found that
253 NTDS was far more efficient for basal area estimation than HPS.

254 While HPS can be a highly efficient system for basal area estimation, the initial
255 investment in training required for proper application of the method is probably greater than for
256 FPS or NTDS. Neither study discloses the prior training level of the technicians who performed
257 the sampling. However, the results may be more easily understood if the technicians had little or
258 no prior experience with HPS. In the Pacific Northwest, where HPS has been widely used for
259 over 50 years, qualified inventory personnel are readily available, and a study employing
260 professional cruisers with multiple years of experience in HPS would perhaps yield very
261 different results.

262 All of the NTDS estimators examined, and the Moore estimator in particular, generally
263 exhibited lower absolute bias with increasing values of n , and therefore the NTDS estimators
264 examined may be more attractive with larger values of n . However, the limited field experience
265 of the primary author in attempting to identify the nearest n trees to a given point suggests that
266 the amount of time required to perform NTDS may increase disproportionately with the value of

267 n desired, particularly when steep terrain or brushy conditions are encountered. Unlike FPS and
268 HPS, where inclusion areas are never mutually exclusive, under NTDS the n tree will always be
269 selected at the expense of the $n+1$ tree. Therefore, a sophisticated recording system or a good
270 memory will be required to efficiently track the distances to all potentially included trees, and
271 this could be challenging with larger values of n .

272 **Conclusion**

273 Of the NTDS estimators examined, only the Moore estimator remained competitive with
274 FPS and HPS at larger sample sizes. The relatively low bias exhibited by the Moore estimator at
275 most sites indicates that it may have potential for estimation of both density and basal area in
276 some forest types. However, hardwood trees exhibiting a clustered spatial distribution (e.g. red
277 alder and bigleaf maple) are common in riparian areas, and our results and those of others
278 (Lessard et al. 1994; Payandeh and Ek 1986) show the Moore estimator to perform poorly for
279 estimating density of clustered populations.

280 Edge-related bias, which can be a particularly problematic issue when sampling long,
281 narrow riparian areas (Lynch 2006), is also of concern. While unbiased correction techniques
282 exist for FPS and HPS (e.g. Ducey et al. 2004), such techniques have only recently been
283 developed for use with NTDS (Lynch *in press*). In riparian forests, which inherently have a high
284 edge-to-area ratio, severe underestimation bias could result from application of NTDS when
285 estimates are not edge-corrected, and therefore the implementation of appropriate correction
286 measures is highly recommended.

287 As sample size increases, the attractiveness of estimation methods which have little or no
288 bias increases relative to estimation methods which have higher bias. This suggests that
289 inventory personnel seeking long-run performance over a large number of sample points will

290 continue to be best served through the use of methods that minimize bias. Although this study
291 has demonstrated that the Moore estimator may have minimal bias in some forest types, it is the
292 opinion of the primary author that none of the NTDS estimators examined (including the Moore
293 estimator) is likely to result in a reduction in measurement costs, relative to FPS and HPS,
294 sufficient to offset uncertainty regarding any insidious bias that may result. The forest inventory
295 community has historically favored design-unbiased estimation methods, and it is recommended
296 that such a preference be retained in this context.

297 Sampling in a highly variable forest type is difficult. Finding too many trees at a sample
298 point, or a string of sample points with no trees, can be psychologically painful, and it is only
299 natural to search for alternatives to this headache. A similar search led some inventory groups to
300 adopt a policy of changing the basal area factor used in HPS so as to get a constant tree count at
301 each sample point (Bell 1994), a policy which has been demonstrated to lead to biased (Wensel
302 et al. 1980; Iles and Wilson 1988) and more variable (Iles and Wilson 1988) results. As an
303 unbiased solution for sample points with too many trees, Iles and Wilson (1988) recommend the
304 plot be split in half, with one side randomly chosen for sampling. This protocol would be useful
305 in FPS as well. There does not appear to be an easy answer for plots with too few trees, but it
306 may be better to accept the added variability and keep the process unbiased.

307 In conclusion, while the ability to control the number of trees sampled for density
308 estimation under NTDS is theoretically attractive, we suggest that none of the NTDS estimators
309 examined will offer operational gains sufficient to offset the relatively poor statistical
310 performance that may result under conditions that surface often in riparian forest sampling.
311 However, the development of new NTDS estimators is currently an area of active research, and
312 future developments may result in estimators that offer sound statistical performance while also

313 being cost-competitive with traditional methods. An estimator developed by Nothdurft et al.
 314 (2010), which requires stem mapping of all n sampled trees, holds particular promise for
 315 sampling the clustered spatial patterns characteristic of riparian areas.

316 **Appendix**

317 The Moore (1954) estimator applies an $(n-1)/n$ multiplier to mitigate the overestimation
 318 bias of the uncorrected NTDS estimator. The sample-point-based estimators for density (trees/ac)
 319 and basal area (ft^2/ac) are, respectively:

$$320 \quad N_M = \frac{n-1}{A_p} \quad ; \quad G_M = \frac{n-1}{n} \sum_{t=1}^n \left[\frac{g_t}{A_p} \right]$$

321 where n is the number of trees to be captured at the sample point, $A_p = \pi(d_n^2)/43,560$ is the area,
 322 in ac, of a circle with radius d_n , d_n is the distance to the n tree, and g_t is the basal area of tree t .

323 Under the Prodan (1968) estimator, the n tree is considered borderline and counted as a
 324 half-tree. The sample-point-based estimators for density and basal area are:

$$325 \quad N_p = \frac{n-0.5}{A_p} \quad ; \quad G_p = \frac{\sum_{t=1}^n [g_t] - 0.5g_n}{A_p}$$

326 Kleinn and Vilčko (2006a) developed an approach based on the arithmetic average of the
 327 distances to the n and $n+1$ trees. Since the distance to the n tree would result in systematic
 328 overestimation and the distance to the $n+1$ tree would result in systematic underestimation, they
 329 reasoned that using the average distance would result in more reasonable estimates. The sample-
 330 point-based estimators for density and basal area are:

$$331 \quad N_K = \frac{n}{A_m} \quad ; \quad G_K = \frac{\sum_{t=1}^n g_t}{A_m}$$

332 where

333

$$A_m = \frac{\pi[(d_n + d_{n+1}) * 0.5]^2}{43,560}$$

334 where d_n is the distance to the n tree, and d_{n+1} is the distance to the $n+1$ tree.

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409

410 Table 1: Density and basal area of each site by tree type.

Site	Conifers		Hardwoods		Total	
	Density	Basal area	Density	Basal area	Density	Basal area
	(trees/ac)	(ft ² /ac)	(trees/ac)	(ft ² /ac)	(trees/ac)	(ft ² /ac)
BL13	111	168	21	13	132	181
KM17	158	286	15	12	173	298
KM18	246	245	15	11	261	256
KM19	215	217	5	4	221	221
KM21	136	197	17	15	153	212
OM36	150	152	15	11	165	163
TH46	167	260	0	0	167	260
TH75	194	244	74	27	268	271

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412

413 Table 2: Clark-Evans (CE) index values by species and site. The CE index takes on values of: 0
 414 if the spatial distribution is extremely aggregated; 1 if the spatial distribution is completely
 415 random; and 2.14 if the spatial distribution is extremely uniform.

Species	BL13	KM17	KM18	KM19	KM21	OM36	TH46	TH75
Douglas-fir	1.03	0.8	1.05	0.93	0.8	1.22	1.11	1.12
western hemlock	-	1.08	0.97	0.9	0.75	-	0.84	-
western redcedar	-	-	0.81	0.82	0.86	-	-	0.48
red alder	-	0.36	0.37	-	0.54	-	-	0.24
bigleaf maple	0.21	-	-	-	-	-	-	0.58
all combined	0.99	1.22	1.09	0.96	0.99	1.25	1.14	1.08

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418 Table 3. Fixed plot sampling (FPS) plot sizes, horizontal point sampling (HPS) basal area
 419 factors, and maximum inclusion areas under HPS, by site, for $n = 6$.

Site	FPS Plot Size (ac)	HPS Basal Area Factor (ft ² /ac)	HPS Max. Inclusion Area (ac)
BL13	0.045	30	0.244
KM17	0.035	50	0.100
KM18	0.023	43	0.111
KM19	0.027	37	0.133
KM21	0.039	35	0.114
OM36	0.036	27	0.209
TH46	0.036	43	0.088
TH75	0.022	45	0.080

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