

Stage duration estimation for *Calanus* populations, a modelling study

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ABSTRACT: Population dynamics of *Calanus finmarchicus* have been modelled using very finely divided representation of the stock according to age-within-stage, in the manner of models developed by C. S. Davis, A. Sciandra, F. Carlotti and others. A key assumption of the model is that development rate is relatively insensitive to food-limitation, so that stage duration can be represented by a temperature function alone. We used the Belehrádek function for this purpose, noting that better data are needed for fitting its parameters. The model closely simulates the timing of stage progression and relative stage abundances of *C. finmarchicus* in the Malangen fjord system (northern Norway) during the winter–spring generation. The model is sensitive to the resolution of the age-within-stage division, but it is fully stable at 0.5 h increments. Modifications of the model simulated several methods for field estimation of stage duration in *Calanus* (or other highly seasonal copepod populations). A method based on changes in stage proportions (the 'Heinle graph' method) is biased by confounding of the effects of developmental progress and mortality on stage proportions. However, the model shows that the bias is mild and the method gives useful estimates of stage duration. Simulation of a method based on molting rate determinations ('Kimmerer experiments') showed its unsuitability for highly seasonal stocks in which stage composition is changing rapidly. Differences in *C. finmarchicus* survivorship schedules between constant and continuously increasing temperatures were simulated, showing that such differences in pattern are critical to annual survival and stock production. Simple methods for fitting mortality rates to data using the model were extremely sensitive to sampling noise. More complex methods may succeed but remain to be developed.

KEY WORDS: *Calanus* · Population model · Stage duration

INTRODUCTION

International planning is in progress for intensive studies of the population biology of marine zooplankton. In particular, the U.S.-initiated GLOBEC program has begun a several-year study of the dominant winter-breeding copepods of Georges Bank and the Gulf of Maine, USA. These are *Calanus finmarchicus* and several species of *Pseudocalanus* (U.S. GLOBEC 1992). The goal is to understand the effects of physical habitat factors on the transfer of nutrition through these copepod populations to stocks of commercially important fish species, particularly cod *Gadus morhua* and haddock *Melanogrammus aeglefinus*. Direct feeding by these fish on copepods occurs

mostly as larvae eating nauplii (Kane 1984, Buckley & Lough 1987). Therefore, estimation of naupliar numbers and explanation of observed variability will be important to the study. Knowledge of the overall dynamics of the copepod populations will be useful for explaining the response of both copepods and their fish predators to variations in marine climatic conditions. Long-term studies of *C. finmarchicus* in Norwegian waters during both the PRO MARE and the ongoing MARE NOR programs underscore the importance of this species in high latitude waters (Tande 1991, Tande & Slagstad 1992).

The dynamics of copepod populations are obscured by great variability in estimates of their numbers. However, if approximately the same progression of

stages is occurring over a reasonably broad stretch of ocean, then it may be possible to derive considerable information about development rates, growth rates, and survivorship from recurring evaluations of stage proportions and body sizes and from time series of field experiments on egg production and molting rates. We have conducted a series of modelling studies intended to show some of the possibilities and limitations of these methods as applied to *Calanus finmarchicus*. The input and verification data for these models come primarily from work in Norwegian fjords, but the implications should apply generally. They lead us to some recommendations for the work on the *C. finmarchicus* stock of Georges Bank and for study of calanid populations everywhere.

Our models follow in some ways that of Carlotti & Sciandra (1989) and Carlotti & Nival (1992) which was developed to describe population cycles in *Euterpina*. Models with similar features were presented earlier by Davis (1984) and Sciandra (1986). The key feature of these models is division of each stage into a number of age steps, with the probability of advance to the next stage increasing from zero only after a minimal stage duration has passed, then rising to 1.0 over an interval mimicking a known or reasonable variability in stage duration. This provides models that are very well behaved with respect to development timing compared to vectorial population models with one abundance element per stage and a single probability of stage advance at each time step (Caswell 1989). Such transition matrix models suffer from unavoidable (and unrealistic) numerical dispersion, so we have abandoned them at the cost of the elegant methods they provide for estimation of stock increase parameters.

We have used our age-within-stage models to study the expectations and biases of stage duration estimates based on (1) recurring estimates of stage proportions and on (2) molting rate experiments as described by Kimmerer & McKinnon (1987) and Peterson et al. (1991). Work with the model showed that the most important parameters for modelling *Calanus* population processes describe the initial, winter spawning pattern, that is, the timing of egg input as the model starts. Data for this can be derived from a suitable time series of egg production estimates (e.g. Diel & Tande 1992). We show some of the implications of the form of this spawning function for stage duration estimates and molting rate experiments. Models of the age-within-stage type can be combined with appropriate field data to approximate survivorship patterns and population production. We spell out the necessary data and discuss the problems of fitting the model and extraction of desired estimates.

LIFE HISTORY AND GROWTH PATTERNS OF *CALANUS FINMARCHICUS*

Calanus finmarchicus stocks spend the late summer and autumn periods in a prolonged diapause that occurs in the fifth copepodite stage (C5). During this rest phase they are deep in the water column, mostly quiescent (although alert and capable of rapid escape responses), and not feeding. Arousal occurs in early to mid-winter, followed by a final molt to the adult stage. Males precede females on average and are available for mating as the females appear. Females (generation G_0 in terminology attributable to Ian McLaren) move upward in the water column and begin to spawn (producing G_1) when abundant food becomes available during the spring increase in phytoplankton. Eggs hatch as nauplii, and there is a progression of 6 naupliar then 5 copepodite stages before return to diapause at depth. In many places there is more than one generation, with the G_2 , G_3 , or even later generations finally entering diapause in late summer. This does not happen in Norwegian fjords (Tande 1991), and it appears that the G_2 generation started by maturation of a fraction of G_1 over the Nova Scotian shelf (McLaren & Corkett 1986) and in shelf water off New England (Miller et al. 1991) never reaches late copepodite stages and probably does not contribute to the following year's stock.

Some uncertainty remains about the relative importance of temperature and food availability in determining the rates of growth and development in calanoid copepods. McLaren (1978) working with Marshall et al.'s (1934) data from Loch Striven has maintained that the timing of development of *Calanus finmarchicus* and *Pseudocalanus* spp. is the same in the field as in the laboratory at the same temperature and at full rations, showing therefore the absence of food limitation. Others have noted that size of individuals (Carlotti et al. 1993) and egg production vary strongly, and that at least egg production (Runge 1985, Nielsen & Richardson 1989, Peterson et al. 1991) appears to be correlated with plant standing stocks as measured by chlorophyll concentration, implying definite food limitation. We contend, based mostly on Miller et al. (1984), that development, the progression from stage to stage, is much less sensitive to food limitation than is growth, the addition of new tissue mass, including egg production. Timing of the stage progression will be relatively insensitive to food availability above a fairly low threshold, while growth, and thus terminal size, will vary as food availability changes. We lack data with which to settle this issue with certainty.

Carlotti & S. Nival (1992) and Carlotti et al. (1993) argue that molting from a given stage will depend upon achievement of a critical weight very close to the terminal weight for the stage. We contend, again

based on the experiments reported by Miller et al. (1984), that hormonal signals for the next molt are received relatively early in each intermolt stage, surely before half of its duration at full nutrition, and that molting proceeds after those signals regardless of subsequently achieved weight. Thus, the physiology of molting tends to sustain the molting rate over a wide range of food availability. Again, the cost of proceeding with a relatively fixed molting schedule, regardless of achieved growth, is variable terminal size. E. G. Durbin & A. G. Durbin (unpubl.) found a strong difference in development of *Calanus finmarchicus* between May 1988 and June 1989 in Great South Channel (Gulf of Maine). In the earlier year progress was at the maximum expected rates for the observed temperatures. In the latter year development slowed and essentially stopped in C4, almost certainly because of very reduced food supplies. Thus, food limitation of development clearly can occur in the field. We only claim that useful modelling results can be achieved by treating food limitation as an unusual circumstance. This allows us to develop our model very simply on the basis of a temperature function for stage duration taken from Corkett et al. (1986; see below).

MODEL FORMULATION

Modelled processes proceed as shown in Fig. 1, a flow diagram; symbols are defined in Table 1. The model is based on recurring evaluation of a series of 12 ($j = 1, \dots, 12$) pairs of vectors, one for each stage from egg through C5. Elements of \mathbf{A}_j in each pair are the abundances of successive ages within the stage; elements of \mathbf{D}_j represent the fraction of the stage duration completed. At each time step all $\mathbf{D}_{j,i}$ are incremented according to the fraction completed in that interval of the stage duration predicted by a modification of the Belehrádek function given by Corkett et al. (1986). This function is

$$\text{Stage duration (d)} = \frac{1}{a_j(T + 10.6)^{-2.05}},$$

where a_j values are given in Table 2. The values 10.6 and -2.05 are parameters fitted to a small set of rearing data. Extended study of *Calanus finmarchicus* populations should include preparation of a much more extensive stage duration data set. As shown in Table 2, we modified the fitted values of a_j so as to approximate growth patterns for *Calanus marshallae* (Peterson 1986) and *Calanus pacificus* (Vidal 1981) which are better studied than that of *C. finmarchicus*. At each time step the conversion is $\mathbf{D}_{j,i} := \mathbf{D}_{j,i} + 1/[a_j(T + 10.6)^{-2.05}]$, where $:=$ means 'replace' in the Pascal language sense. The units of $\mathbf{D}_{j,i}$ are d^{-1} . A copepod living at temperature T will complete stage j in $a_j(T +$

$10.6)^{-2.05}$ days, so in 1 d it completes the fraction $1/[a_j(T + 10.6)^{-2.05}]$ of its development toward the next molt. These fractions can be added for successive days (at varying temperatures) until the sum reaches 1., when molting occurs to stage $j+1$. Thus, $\mathbf{D}_{j,i}$ is the cumulative fractional completion of development for the stage.

Each vector pair is evaluated once on each time step. In the simplest case there is no variation among individuals in stage duration. For each element pair, when the fraction $\mathbf{D}_{j,i} \geq 1.0$, then the abundance estimate $\mathbf{A}_{j,i}$ is transferred to the first element of $\mathbf{A}_{j+1,1}$, that is, $\mathbf{A}_{j+1,1} := \mathbf{A}_{j+1,1} + S_j \cdot \mathbf{A}_{j,i}$, where S_j is the survivorship assigned to stage j . A variable number of elements of \mathbf{A}_j may have $\mathbf{D}_{j,i} \geq 1.$ on any given time step, so that $\mathbf{A}_{j+1,1}$ can fluctuate. In the case of C5's, elements of \mathbf{A}_{12} with $\mathbf{D}_{12,i} \geq 1.$ are added to the new resting stock resulting from G_1 . When the value of $\mathbf{D}_{j,i}$ is less than 1.0, then both the abundance and development fractions are transferred to the next age-within-stage element: $\mathbf{A}_{j,i+1} := S_j \cdot \mathbf{A}_{j,i}$ and $\mathbf{D}_{j,i+1} := \mathbf{D}_{j,i}$. Incrementing of $\mathbf{D}_{j,i}$ and transferring the new value to $\mathbf{D}_{j,i+1}$ are separated in the program sequence. At each time step the first elements of all development vectors are set to zero, $\mathbf{D}_{j,1} := 0.$ for $j = 1$ and 12. In a slightly more complex case, a frequency distribution for molting as a function of $\mathbf{D}_{j,i}$ is applied at each time step. This mimics the effect of variable stage duration.

Stage-specific survivorship rates, S_j , were arbitrarily selected for the standard run so as (1) to allow overall

Table 1. Symbols used in presenting the *Calanus finmarchicus* population model

Symbol	Definition
j	Stage number: 1 = egg, 2 = N1, 8 = C1, 13 = C5-diapause
i	Time step number
$\mathbf{A}_{j,i}$	Abundance of stage j in its i^{th} age class
a_j	Belehrádek constant for duration of stage j
B	Stock biomass (dry weight mass m^{-2})
C1–C6	First to sixth copepodite stages
$\mathbf{D}_{j,i}$	Fraction of stage development completed by individuals of i^{th} age within stage j
JD	'Julian Day', consecutive number of day of the year; JD 1 = 1 January
G_0	Number of generation completing diapause early in model year
G_1	Offspring of G_0 females
g_j	Growth rate of stage j
MR_j	Molting rate of stage j (including all ages within j)
N1–N6	First to sixth naupliar stages
S_j	Finite daily survivorship of stage j
$W_{j,i}$	Weight of stage j individuals at $\mathbf{D}_{j,i}$
$W_{10,j}$	Newly molted weight of stage j individuals

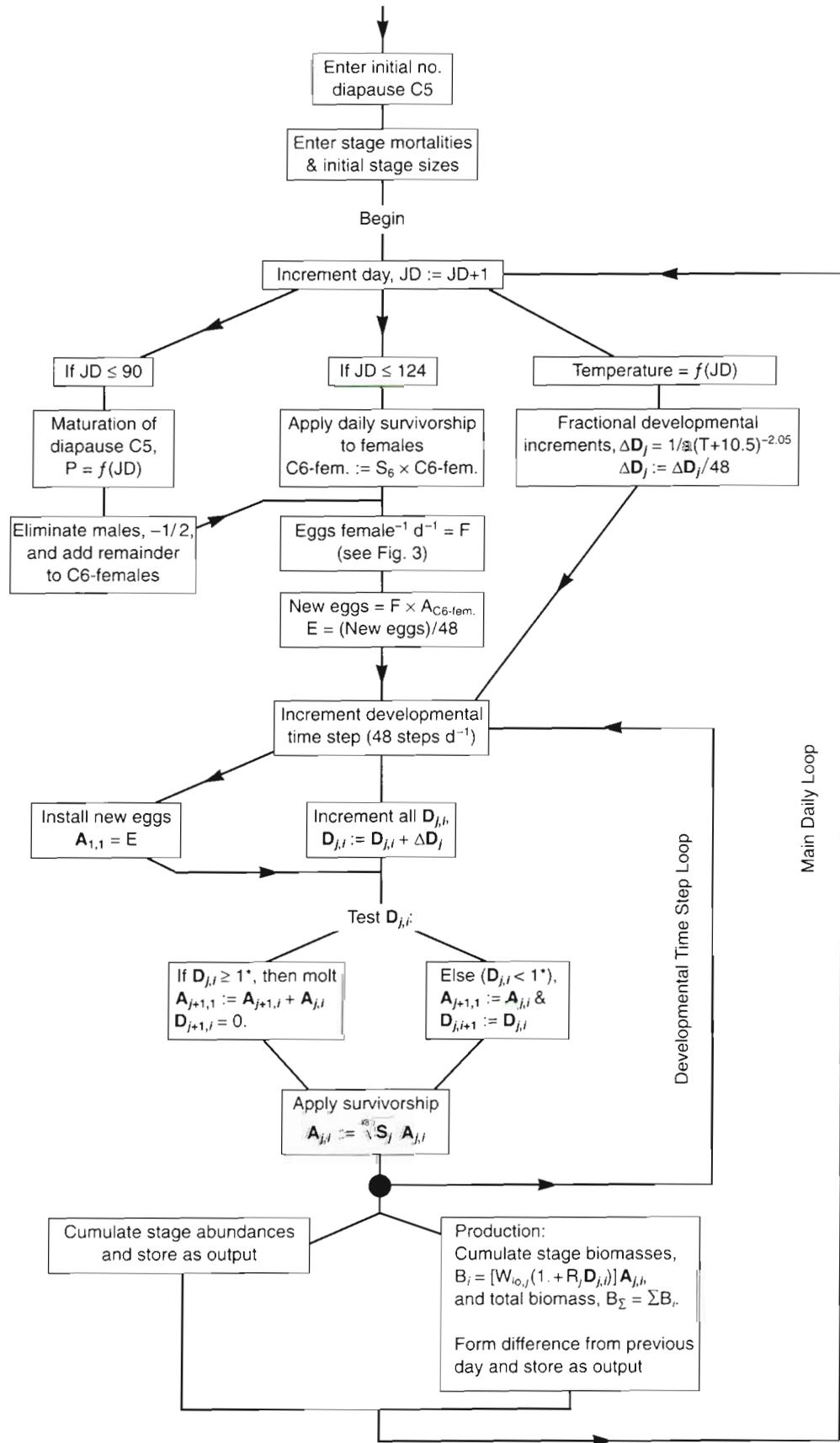


Fig. 1. Process flow diagram for age-within-stage model of population development in *Calanus finmarchicus*. Alternative criteria for variable development timing are applied at the single * (development time step loop). Regression coefficients for growth rates of stages, R_j (applied at bottom right), are given in Table 2

Table 2. Parameters of the MALANGEN version of the *Calanus finmarchicus* population model. JD: Julian Day

Temperature: $T(\text{JD}) = 6.25 + 2.75 \sin[2\pi(199 + \text{JD})/365]$				
Maturation probability for diapause C5: cumulative normal with μ at JD = 45, $\sigma = 20$ d				
Female survivorship after maturation (finite daily rate): 0.985 d^{-1}				
Egg production function: Eggs female ⁻¹ = $27. \times [1. - (107 - \text{JD})/33.]$, for JD = 74 to 107 Eggs female ⁻¹ = $27. \times [1. - (\text{JD} - 107)/17.]$, for JD = 108 to 123)				
Belehrádek function coefficients for stage durations and finite daily survivorship rates for naupliar and copepodite stages:				
Stage	a (coefficient from Corkett et al. 1986)	a (coefficient used in MALANGEN)	Finite daily survivorship	Lowest stage weight, W_{i0} ($\mu\text{g dry wt ind.}^{-1}$)
Egg	691.	691.	0.850	0.700
N1	1069.	1069.	0.920	0.700
N2	.	.	.	0.700
N3	.	.	.	0.700
N4	.	.	.	0.800
N5	.	.	.	0.900
N6	1069.	1069.	0.920	1.000
C1	1595.	1595.	0.950	2.317
C2	1802.	1700.	0.960	4.810
C3	1785.	1810.	0.970	9.990
C4	1925.	1925.	0.980	20.75
C5	3951.	5584.	0.990	43.12
C5-diapause			1.000	89.64 (fixed)
Biomass as a function of stage and age-within-stage:				
N1 and N2: Stock biomass = (Numbers in age step) W_{i0}				
N3 to N6: Stock biomass = (Numbers in age step) $(W_{i0} + 0.1 \mathbf{D}_{j,i})$				
C1 to C4: Stock biomass = (Numbers in age step) $[W_{i0} (1. + 1.077 \mathbf{D}_{j,i})]$				
C5 to diapause: Stock biomass = (Numbers in age step) $[W_{i0} (1. + 2.33 \mathbf{D}_{j,i})]$				

survival of approximately 4 individuals entering diapause as C5 for each hundred eggs spawned, (2) to concentrate mortality in the naupliar phase, and (3) to give 'reasonable' relative abundances of the copepodite stages. Tests of various other schedules of S_j are reported.

The mechanics of the model in these respects are like those described by Carlotti & P. Nival (1992, particularly their Fig. 2). However, we made important changes. They have an 0.25 d step for age-within-stage, but an hourly time step for the model run. In our model the vectors have an age step for each time step. Our model was based on finite differences from the outset, not as an implicit solution to a set of differential equations as in the case of the Carlotti models. Another difference is that those models are driven by submodels that produce data about food availability, which (as discussed above) is not a concern in our present study. Models of the type used here differ from

those employed by Slagstad (1981), Slagstad & Tande (1990), and Tande & Slagstad (1992), where development was a continuous function of temperature and food availability, and stages were arbitrarily assigned to successive sections of the development scale (a model of this general kind is used by Miller in press in a study of *Neocalanus plumchrus*). The effects of this difference have not yet been very fully explored. The Slagstad models are also driven by submodels for food availability.

Our model represents the calendar year from 1 January onwards, shown in our graphs on a 'Julian Days' (JD) axis. It is initiated by specifying the number of resting C5 m⁻² in the region to be modelled, half of which are taken to be male. Resting C5 mature at a fractional daily rate which increases in sigmoid fashion (Fig. 2) from 0 to 1. Once matured, females are subject to a mortality rate which interacts with the maturation schedule to produce the time course of female abundance. On each model date the per capita egg output rate is set by a function (Fig. 3) derived from the data collected from the *Calanus finmarchicus* population in Malangen fjord in 1989 (Diel & Tande 1992). Several other spawning rate functions of different forms were tried to show the influence of this initiating function on the progression

of stages in the model. Each day's total egg production (no. of females \times per capita daily fecundity) is allotted into the vector of egg abundance according to the number of time steps per day. Thus, if the time step is 0.5 h, then 1/48 of the total eggs for the day are installed at each step in $\mathbf{A}_{1,1}$, the first age-within-stage element of the egg vector.

The model provides a rough estimate of biomass production attributable to the growth of G_1 . Each model day the biomass is calculated by multiplying the stage abundances (\mathbf{A}_j) by an individual weight (W_j , mg dry wt) appropriate to the developmental progress within the stage ($\mathbf{D}_{j,i}$). These products are summed over the stages ($B = \sum_j \sum_i W_j \cdot \mathbf{A}_{j,i}$), and productivity is the increment over the previous day's biomass. Weight increments within stages are added to estimates of the initial weights of the stages, W_j , taken from Peterson's (1986) data for *Calanus marshallae*. This Pacific species has the same range of terminal weight as

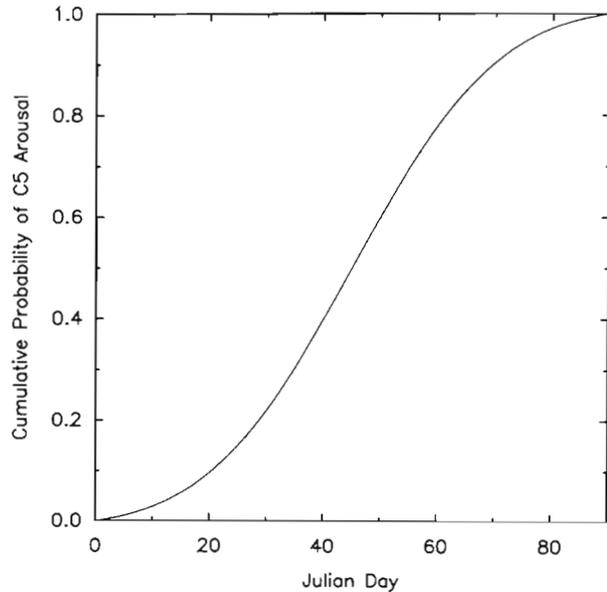


Fig. 2. Probability of individual C5 breaking diapause and maturing, as a function of date

C. finmarchicus, and Peterson's data provide careful weighings for all stages. Weight increments were calculated from the slopes of Peterson's regressions of weight vs stage (Table 2). Carlotti et al. (1993) have provided an analysis of the effect of temperature on body weight as a function of habitat temperature in *C. finmarchicus*. Their function could be incorporated

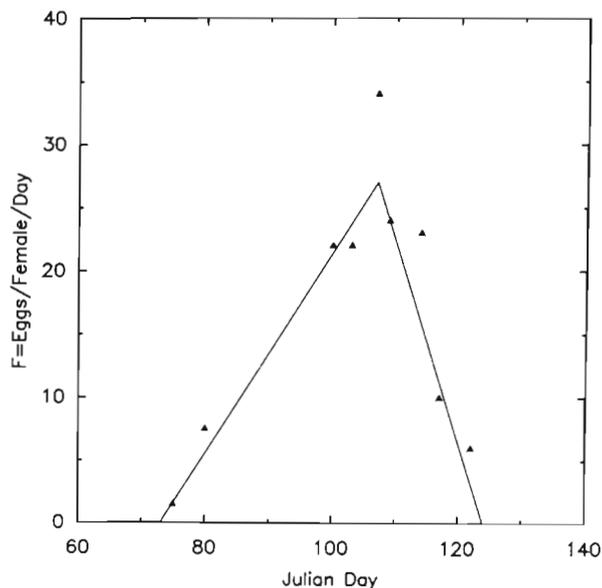


Fig. 3. Per capita egg production rate in Malangen Fjord from data of Diel & Tande (1992). Plotted points are the actual estimates; fitted lines are the spawning function used in MALANGEN. Overall rate varies with changes in both eggs per spawning and spawning frequency

in the model to give a more precise estimate of biomass production, but for now that has not been done.

Model temperature is specified by a simple sine function. This was varied with the location to be simulated, and the approximation is shown for each case presented. Results are presented showing the impact of different temperature patterns.

The realization of the model is a FORTRAN-77 program which was run on a Sun SPARC-1+ workstation. Copies of the program for our standard run or for special 'experiments' are available from the authors on request.

SPECIFIC SIMULATIONS, RESULTS AND EXPLANATIONS

Our results divide into 2 groups: studies of the characteristics of the model itself, and simulations of techniques for evaluation of developmental progress (stage duration) in the field.

The Malangen case, a standard run

Parameter values employed in our standard version of the model (hereafter called MALANGEN) are given in Table 2. All of these values are intended to represent the population processes of *Calanus finmarchicus* in Malangen Fjord, northern Norway. Seasonal temperature variation in the model (Fig. 4) was fit to data from 20 m in Malangen Fjord from February to December 1990. Interaction of C5 maturation (Fig. 2), female mortality (0.015 d^{-1}), and the spawning function (Fig. 3) produced a total of 1.03×10^6 eggs from an original stock of ca 5000 females, an average just over 200 eggs apiece, well within the production capacity of the species, but reasonable since mortality removes many before spawning is complete. G_0 females were arbitrarily removed after the spawning function dropped to zero by increasing their mortality to 0.15 d^{-1} .

Survivorship values were selected to give about 4 G_1 -C5 entering diapause at the end of the growing season for each G_0 -C5 maturing after 1 January. Overall survivorship of 4% allows for an 0.0067 d^{-1} mortality (6.68 deaths per thousand individuals per day) over the 210 d diapause phase, bringing the G_1 stock back to the starting place of G_0 , as it becomes the new G_0 the following winter. We have no idea whether this is a reasonable mortality rate for the resting stock or not, but new field research could give an approximate answer. The survivorship value for nauplii produces a peak naupliar abundance of 13 l^{-1} (integrated over 25 m) (Fig. 5), in general agreement with observations

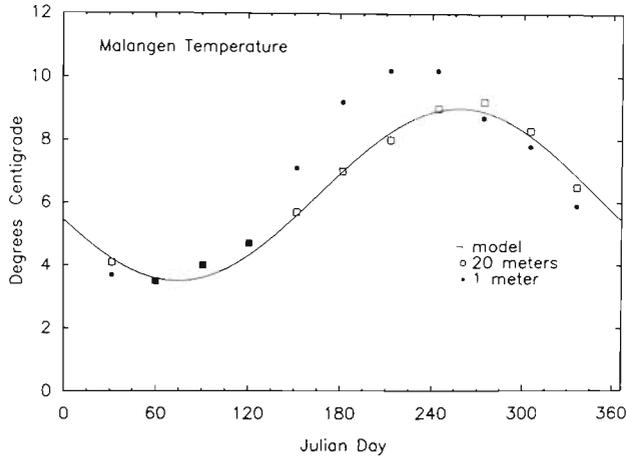


Fig. 4. Solid curve: temperature variation in Malangen Fjord from a fitted sine function. Open boxes show 20 m temperature in Malangen Fjord, 1990. Filled boxes show that surface layer was slightly warmer in spring and summer. *Calanus finmarchicus* avoids this layer because of reduced salinity

from Lofoten, Norway. In an elaborate study by Ellertsen et al. (1989) of the spawning behavior and larval survival of cod, the abundance of *Calanus* nauplii peaked in April at 15 to 20 nauplii l^{-1} in 4 yr between 1980 and 1985. MALANGEN first produced copepodites (Fig. 5) on JD 105 (April 15), and the last C5 entered diapause on JD 185 (July 4). The timing of the stage progression is realistic for north Norwegian fjords, and copepodite abundances are within the range observed in fjords and shelf areas during the annual recruitment (Tande & Slagstad 1992). Survivorship and mortality curves (Fig. 6) for the G_1 stock show concentration of mortality in the naupliar stages. This

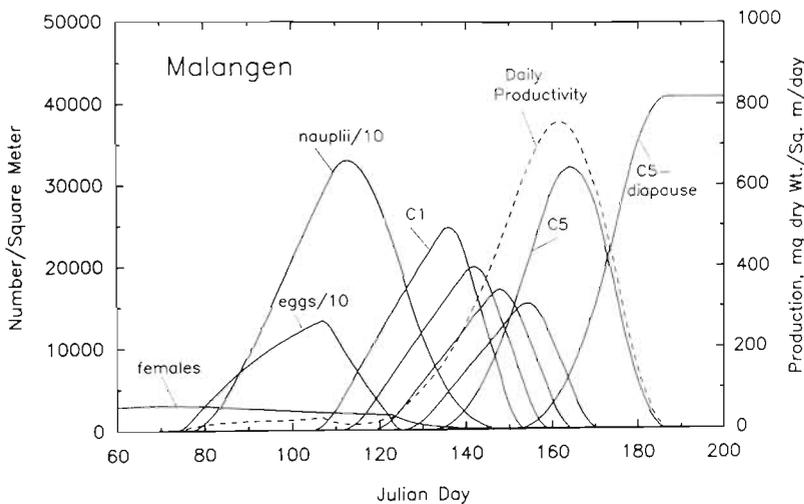


Fig. 5. Output of the standard or MALANGEN version of the model. Abundances of successive stages shown by solid lines. Daily production increments shown by dashed line. Total production was $12.3 \text{ g dry wt m}^{-2}$. Initial number of G_0 C5 was $10\,000 \text{ m}^{-2}$

mostly occurs simply because the naupliar phase comes first and lasts nearly half the life cycle. There are more of them, so they suffer the most mortality regardless of small differences in rates between stages.

Effect of time step size

MALANGEN has 48 ($\frac{1}{2}$ h) time steps per day. The choice of time-age steps has strong effects on the model. A run like MALANGEN, but with just one time and age step per day (Fig. 7), has substantially lower survival (2.9 G_1 -C5 entering diapause per hundred eggs spawned) and production (annual total = 8.83 g dry wt) than runs with higher resolution. Both survival and production are nearly asymptotic with half-hour time steps, being very slightly less than from a run with 200 time-age steps d^{-1} . The reason for this improved survivorship is that development is significantly faster. All 'individuals' (all $D_{j,i}$) that reach 1.0 can graduate nearly immediately, rather than having to wait for the end of the day. For example, if the required stage duration is 5.1 d, the 'saving' of time in a high resolution model will be 0.9 d. Moreover, since mortality is modelled as decreasing with advancing stage, overall survivorship is also enhanced by reduction of time at younger, more susceptible stages. It is possible that a model might be desired in which molting was restricted to some part of the day, say night (Miller et al. 1984). In that case it is better to explicitly add this condition on molting to the model, rather than to let long time steps provide the effect artificially. In addition to enhanced survivorship, the output of high

resolution models is smooth at the resolution of the plotting, which is not true at one step per day (Fig. 7). The drawback of high resolution is running time. At 1 time-age step d^{-1} , MALANGEN runs in under 2 s (Sun SPARC 1+), whereas at 48 steps d^{-1} it requires over 5 min.

Incorporation of refined time steps must be done with care about the form of the development and survivorship functions. The daily increments to $D_{j,i}$ must be divided by 48 for application each half hour. As we have formulated survivorship, it is entered as a finite daily rate, which with one time step per day is multiplied by the abundance to give the new abundance. Thus, $A_{j,i+1}(JD+1) := S_j \cdot A_{j,i}(JD)$. If survivorships are to be applied twice daily by this multiplicative rule, giving the

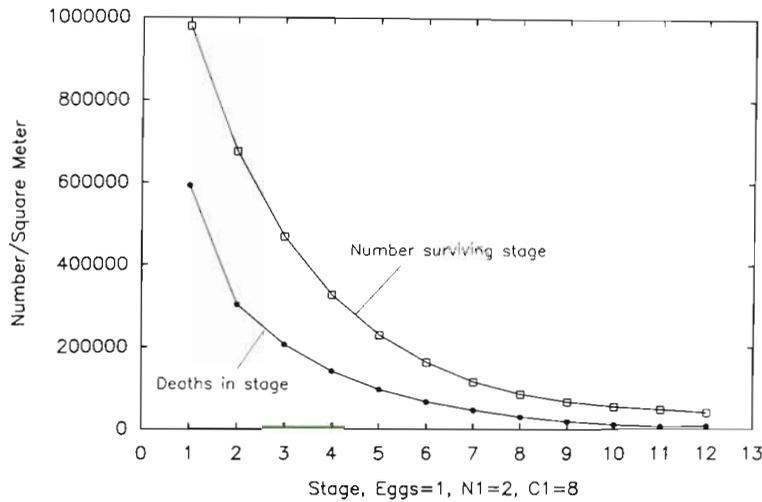


Fig. 6. Stage-by-stage survivorship and mortality curves for the MALANGEN model

same result, the replacement becomes $A_{j,i+1}(JD+\frac{1}{2}) := (S_j)^{1/2} \cdot A_{j,i}(JD)$. For half-hour time steps, the survivorship multiplier is the 48th root of S_j .

All versions of the model, including MALANGEN, have a good deal of hidden numerical 'chatter'. This is brought out by plotting the age-within-stage distribution of a given stage for several dates during the passage of the stock through that stage (Fig. 8). These are plots of successive $A_{10,i}$ (C3) for all occupied age-within-stage steps. The strong, narrow spikes develop because several $A_{j,i}$ can transfer to $A_{j+1,1}$ on some time steps. The numbers of elements

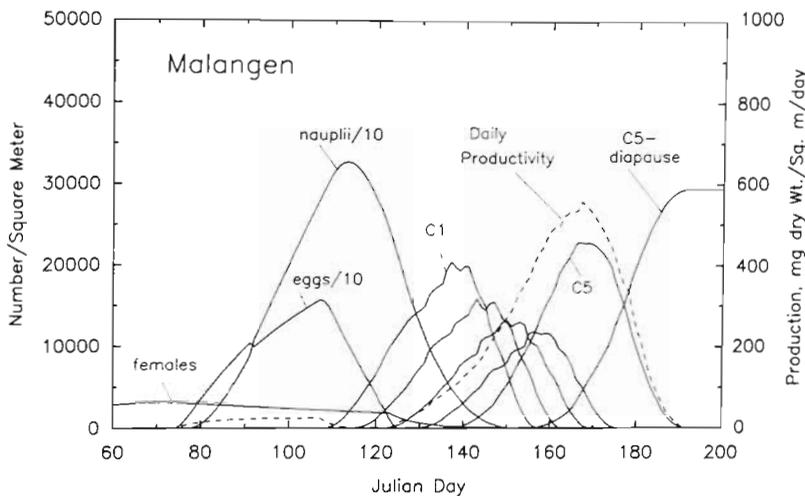


Fig. 7. Output of model with same parameters as MALANGEN, but only one time step per day. Stage abundance sequences are somewhat erratic. Both survivorship and production (total = 8.83 g dry wt m⁻²) are reduced. The 'tooth' in egg abundance at JD = 90 is caused by completion on that day of maturation of G₀-C5's. For 2 d female mortality overrides the increasing per capita fecundity (Fig. 3)

$A_{j,i}$ making this transfer and their total may vary widely among time steps. So far as we can determine, this numerical chatter does not impede the reliability of the model as a representation (Fig. 5) of development timing in *Calanus finmarchicus*.

Effects of variable stage duration

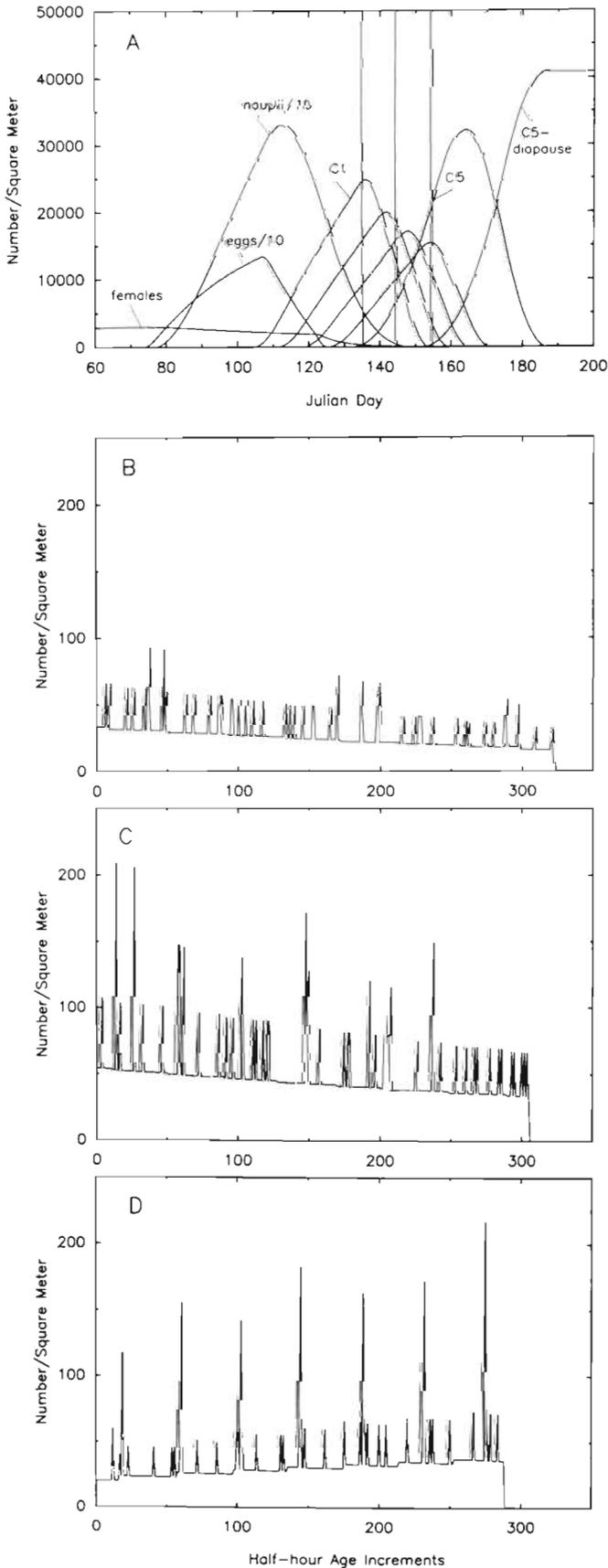
Five simulations were run to examine the effects of variation of stage duration between individuals on timing of the stage progression and on other aspects of model output. In VAR-1 through VAR-5 the probability of molting was allowed to increase as a function of $D_{j,i}$ in the patterns shown in Fig. 9. All patterns that allow some individuals to molt early

(VAR-1, VAR-2, VAR-4) gave enhanced overall survival and productivity together with completion of later stages several days ahead of the MALANGEN schedule (Table 3). Enhanced survival in those cases derived from earlier arrival of some individuals at later stages, reducing the duration of their exposure to the higher mortality rates for early stages. Patterns in which some individuals had their molting delayed to various degrees (VAR-3 and VAR-5) were only slightly affected, with small reductions in survivorship and production. The impact of early advance was much stronger than roughly equivalent delay (VAR-1 and VAR-2) when

both occurred together. The model suggests strong selective advantage will accrue to individuals that reduce stage duration. The effect is such that average stage durations we observe in *Calanus finmarchicus* today must already have been minimized by recurring selection of this kind in the past. There must also be countervailing selective costs preventing further shortening. A different sort of model has been used by Myers & Runge (1983, 1986) to explore the selective tradeoffs among growth rate, stage duration, and survivorship.

Effects of the temperature increase schedule

Timing of the seasonal temperature cycle has a strong effect on the model and presumably on the timing of real



Calanus finmarchicus cohorts. As a simple example, MALANGEN was modified only by changing the temperature function to fit data from Georges Bank (Fig. 10A, where the seasonal low is comparable to that in fjords in northern Norway, but the spring warming is more rapid and the summer maximum higher. Overall survivorship (Fig. 10B) increased to 64 000 ind. m⁻² and production increased by over 50%. Copepodites appeared at nearly the same time, but C4 was completed 12 d earlier and C5 16 d earlier. The greater rapidity of development (enforced, of course, by the Belehrádek function) accounts for the improved survivorship and production through reduction of exposure to mortality at fixed rates. Reality of a Georges Bank version eventually can be enhanced by including a site-appropriate spawning function, although the temperature difference from Malangen isn't great during the spawning season. Mortality rates may increase at higher temperatures through a general speeding up of all ecological processes including predation.

Sensitivity analyses have been carried out to examine the consequences for a field population of the effects of a temperature regime studied experimentally by Pedersen & Tande (1992). In their experiments, eggs and then nauplii were held at 2°C until completion of the late naupliar phase, then programmable temperature control baths held one group at 2°C, warmed another at 0.1°C d⁻¹, and warmed a third at 0.2°C d⁻¹. Mortality observed in the laboratory was progressively less severe for those groups given faster warming. In order to look at this aspect in more detail, we performed 2 series of model tests (Table 4). In the first series (PT-1 to PT-5) MALANGEN was changed only with respect to temperature. It was held constant until JD = 120, then increased according to the schedules in Table 4 (Series 1). Both integrated survivorship and production increased monotonically with the rate of temperature increase because of reduced stage duration and, thus, shorter exposure to fixed mortality rates. However, all simulations had less productivity than MALANGEN because of lower temperatures early in the season and later onset of warming. Not only the rate of increase (actually higher in PT-5 than in MALANGEN), but also the time when warming starts in spring affects the final outcome. Delay in warming slowed arrival of the stock at early copepodite stages (10 d delay of the C1 peak in PT-4 compared to

Fig. 8. Age-within-stage distributions for C3 from MALANGEN. Age distributions on 3 dates indicated by vertical lines in (A) are shown by half hour intervals in (B) for JD = 135, in (C) for JD = 145 and in (D) for JD = 155

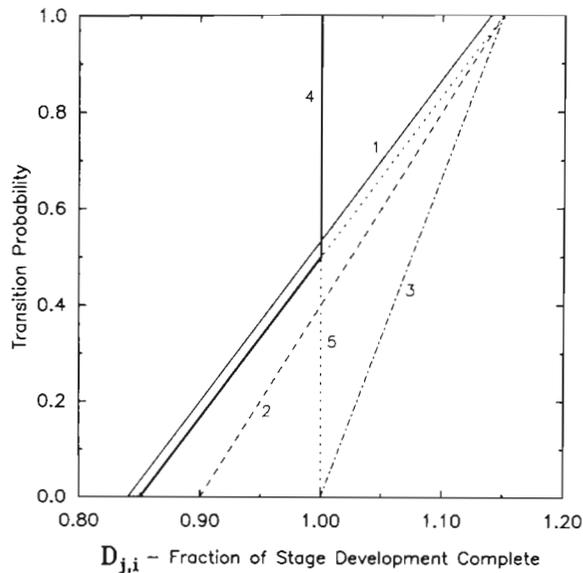


Fig. 9. Patterns of variation in probability of molting, P_m , as a function of $D_{j,i}$. Patterns were (1) linear increase from $D_{j,i} = 0.85$ to 1.15; (2) linear increase from $D_{j,i} = 0.90$ to 1.15; (3) linear increase from $D_{j,i} = 1.0$ to 1.15; (4) linear increase at same rate as pattern (1) from $D_{j,i} = 0.85$ to 1.0, then $P_m = 1.0$; (5) $P_m = 0$, until $D_{j,i} = 1.0$, then linear increase at rate of pattern (1)

MALANGEN) with a tendency to more overlap in successive abundance curves later in the spring.

In the second series (PT-6 to PT-9), we evaluated the effect of variable survivorship regimes derived from the instantaneous mortality constants determined in laboratory rearings (Pedersen & Tande 1992). These alternate survivorship schedules were initiated on model day JD = 120, the approximate equivalent of the experimental transfer of experimental stocks to different increase regimes. In the experiments, the largest proportions of the stock when different warming schedules were applied were in N4 and N5, roughly the same as for MALANGEN on JD = 120. Although the naupliar development period found by Pedersen & Tande was longer than expected for 2 °C according to Corkett et al. (1986), the discrepancy (for

which reasons were given by Pedersen & Tande) does not affect the conclusions from our model. We adapted the survivorship schedules from Pedersen & Tande (1992, their Fig. 3) as shown in Table 4 (Series 2) and Fig. 11. After JD 120 in PT-6 to PT-9 there was no differential in survivorship between stages. Development was driven by the MALANGEN warming sequence; the mortality variations among PT-6 to PT-9 were not compounded by the slowing of development resulting from slower warming. The simulations indicated that the non-predatory mortality observed in the absence, or near absence, of warming had severe impact on final cohort survival, and thus on productivity. In PT-6 no C5 reached diapause, and in PT-7 it was only a very few. High daily survivorship (PT-9), as obtained with the fastest experimental warming, produces unrealistically high survivorship and production. It shows, however, the favorable impact of warming on the stock, apart from possible effects on predation. The subtle step in survivorship from PT-7 to PT-8 shows that the model is very sensitive to variations in survivorship constants. The result underlines the contention (Pedersen & Tande 1992) that substantial changes in productivity are likely to occur if a temperature-dependent mortality operates during the annual recruitment period in *Calanus finmarchicus*. It clearly appears that more data on field mortality are needed in order to further explore this subject and substantiate this mechanism.

Stage duration derived from time series of stage proportion estimates

Two approaches to determination of stage duration in the field have been simulated with the model. We call these the 'Heinle graph' and 'Kimmerer experiment' methods after their originators. Neither method was developed for precisely the problem to which we apply it. Heinle graphs were originally applied to determination of stage durations in laboratory experiments (Heinle 1966, Johnson 1981, Landry 1983, Peterson 1986, Peterson & Painting 1990). They have also been applied by Miller & Nielsen (1988), and Miller (in press) to field populations of *Neocalanus plumchrus* sampled in the Gulf of Alaska, and by Durbin & Durbin (unpubl.) to *Calanus finmarchicus* in Great South Channel (Gulf of Maine). They are constructed by collecting a time series of samples of the population and determining the proportions of the stages in each. Cumulative values of the proportions at each sampling time

Table 3. Output of MALANGEN modified to use variable molting functions as shown in Fig. 9

Molting probability pattern	Overall survivorship to C5-diapause (thousands)	Integrated production (dry tissue weight, g)	JD at peak of C4 stock
MALANGEN	41	12.3	154
VAR-1	58	17.5	149
VAR-2	50	15.0	151
VAR-3	37	11.2	157
VAR-4	58	17.5	149
VAR-5	40.5	12.1	154

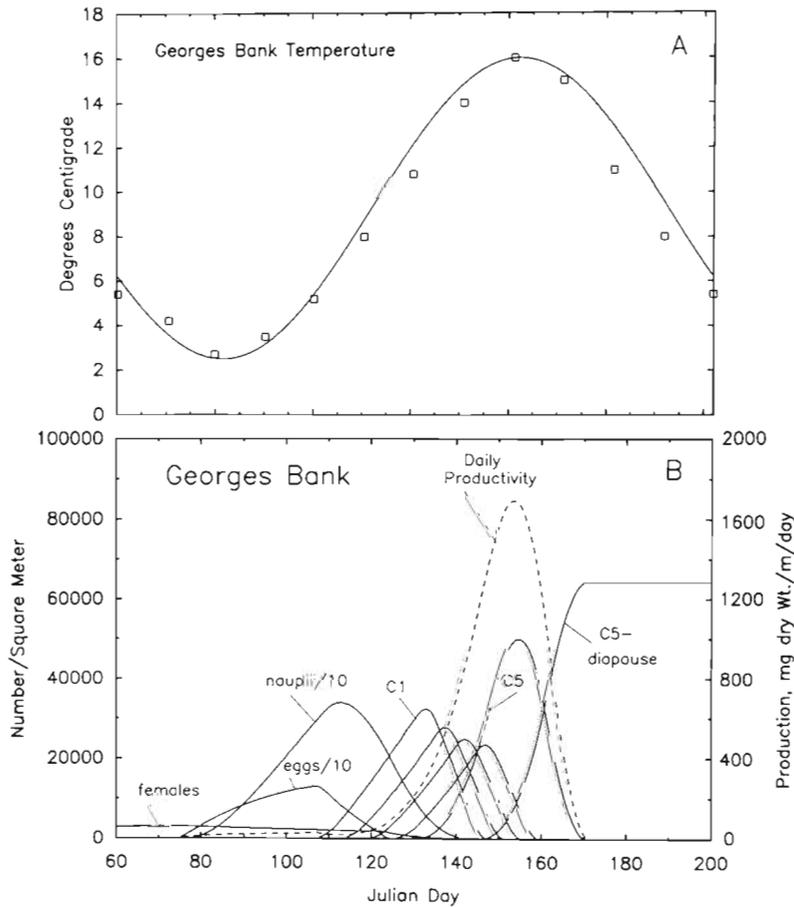


Fig. 10. (A) Temperature function for Georges Bank based on surface temperature data (open boxes) from Walsh et al. (1987). (B) Output of Georges Bank *Calanus finmarchicus* population model. Only the temperature function was changed from MALANGEN. Total production was 19.3 g dry wt m⁻²

Table 4. Alternate mortality schedules for model versions exploring the consequences of the Pedersen & Tande (1992) rearing results

Series	Simulation no.	Temperature (°C)	Survivorship rate schedule	Productivity (g dry wt m ⁻²)
1	PT-1	2.0	As MALANGEN	2.53
	PT-2	2 + 0.05 d ⁻¹	As MALANGEN	4.32
	PT-3	2 + 0.10 d ⁻¹	As MALANGEN	6.04
	PT-4	2 + 0.15 d ⁻¹	As MALANGEN	7.69
	PT-5	2 + 0.20 d ⁻¹	As MALANGEN	9.15
2	PT-6	As MALANGEN	As MALANGEN to JD = 120, then as Fig. 10	0.44
	PT-7	As MALANGEN	As MALANGEN to JD = 120, then as Fig. 10	0.95
	PT-8	As MALANGEN	As MALANGEN to JD = 120, then as Fig. 10	14.11
	PT-9	As MALANGEN	As MALANGEN to JD = 120, then as Fig. 10	40.8

are then plotted as a vertical series, and points representing the sequence of cumulative proportions of each stage are joined by lines or represented by regressions against time. These 'transition curves' delineate the progress of the cohort's transition between successive stages.

The Heinle graph (Fig. 12A) for MALANGEN was calculated using data for only copepodite stages, leaving out nauplii. We believe this represents a realistic situation for field work, in which naupliar and copepodite stages must be estimated in different ways such that comparable numbers would be very difficult to obtain. A problem with this representation is that it does not include all of the cohort at the outset, only those individuals that have reached C1. Thus, the form of the early transition curves is greatly influenced by continuing recruitment from the naupliar stock. However, the form of the C3→C4 and C4→C5 transition curves is only very slightly influenced by continued recruitment to C1. In field work a check on the validity of this assumption could be obtained by examining the abundance of late nauplii. If it is negligible by the date at which C3→C4 crosses the median (50% of total stock), then recruitment bias should be small. The time elapsed between crossings of the median line by the C3→C4 and C4→C5 transition curves is an only slightly biased estimate of the C4 stage duration. Precisely, the estimate for MALANGEN was 5.83 d (indicated by vertical lines in Fig. 12A), while the Belehrádek function expectation for the 5.66°C temperature at the midpoint on JD = 153.47 was 6.33 d. The downward bias is caused by the differential in mortality among the stages (Hairston & Twombly 1985, Miller in press). If mortality is progressively less at successive stages, then later stages will accumulate *relative* abundance faster than their development rate alone would predict, shortening the time elapsed between crossings of the median by successive transition curves. For MALANGEN the bias is -0.50 d, or 8%.

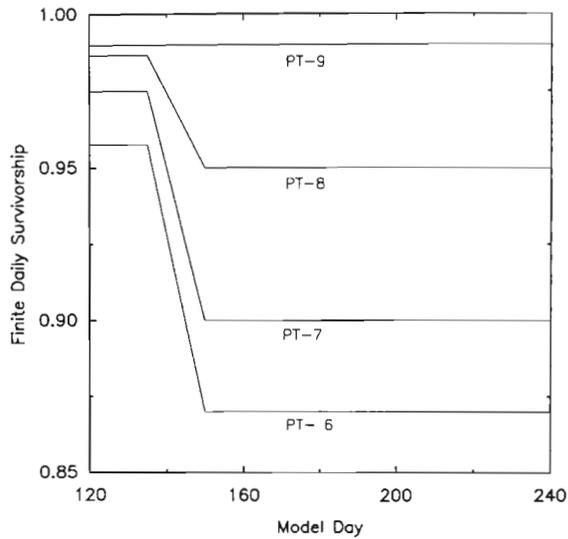


Fig. 11. Survivorship patterns applied after JD = 120 in PT-6 to PT-9. These patterns represent the effects on survivorship observed for different warming schedules by Pedersen & Tande (1992). PT-6 was observed at constant 2°C , PT-7 with warming at $0.05^{\circ}\text{C d}^{-1}$, PT-8 with warming at $0.15^{\circ}\text{C d}^{-1}$, and PT-9 with warming at $0.20^{\circ}\text{C d}^{-1}$

Detailed simulations of varied mortality differentials show that this source of bias operates in our present models exactly as a different model by Miller (in press) suggested it should. That is, bias arises only from differentials in mortality between stages, not from mortality per se. A variant of MALANGEN called CONSTSURV was developed by making survivorship rates identical among stages, then changing its value until overall survivorship (Fig. 13A) was the same (output was 40851 C5 m^{-2} in diapause) as for MALANGEN. The Heinle graph estimate of C4 duration from CONSTSURV (Fig. 12B) was within 0.1% of the Belehrádek function expectation of 6.28 d. The very slight discrepancy was due to nonlinearity in the time course of temperature change. Next, variants were developed with survivorship increasing $2\% \text{ d}^{-1} \text{ stage}^{-1}$ through the copepodite stages (INCSURV, Fig. 13B) and decreasing $2\% \text{ d}^{-1} \text{ stage}^{-1}$ (DECSURV, Fig. 13C). INCSURV required some increase of naupliar survivorship (Fig. 13B) to provide enough copepodites to support the high mortality in early copepodite stages. DECSURV required extremely high naupliar survival to provide enough copepodites to support the prolonged mortality at high rates in late

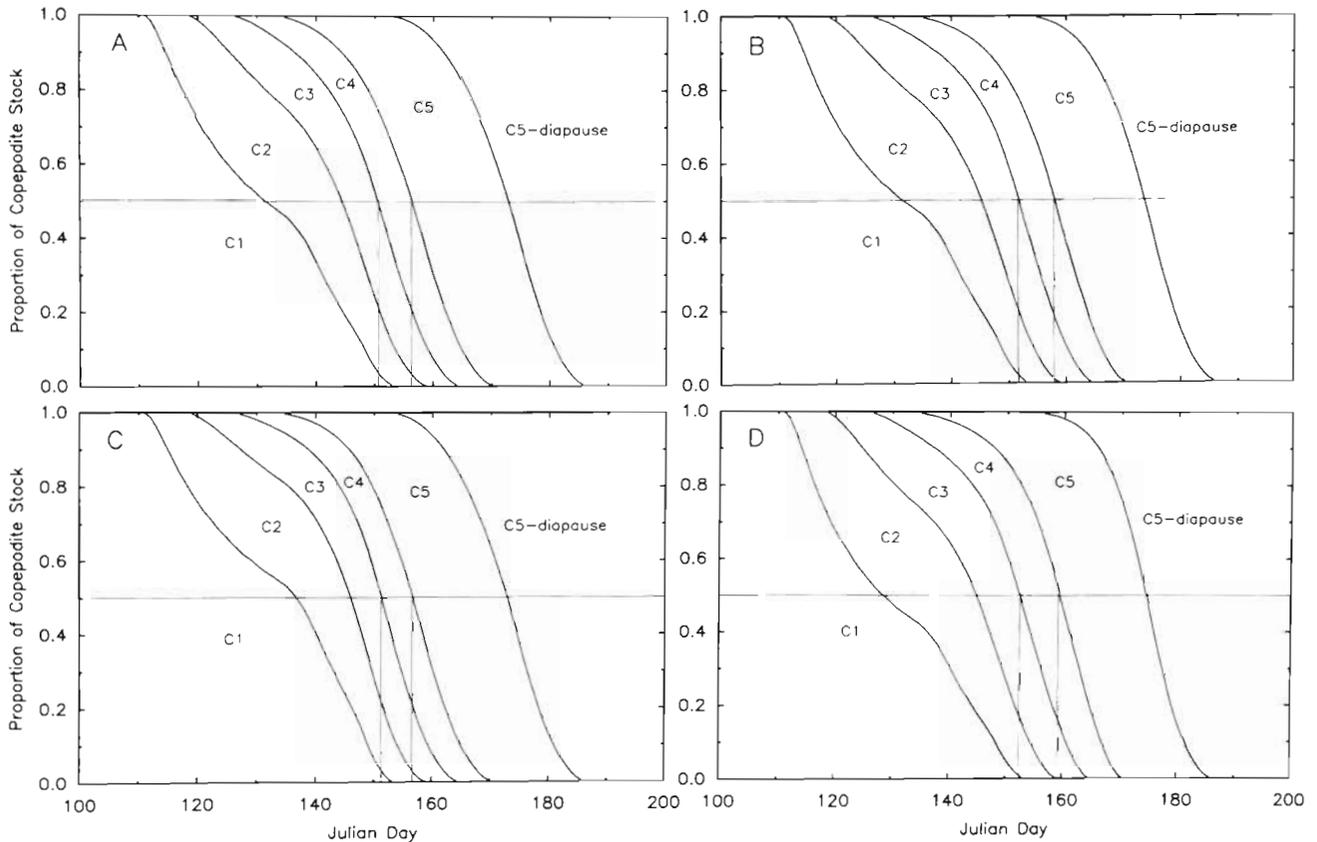


Fig. 12. (A) Stage proportion or Heinle graph for MALANGEN. Calculation of stage proportions is for copepodite stages only. Apparent duration of C1 and C2 is extended by continued recruitment from naupliar phase. Apparent duration of C4 is shown by vertical bars. (B) Heinle graph for CONSTSURV. (C) Heinle graph for INCSURV (copepodite survivorship increased $2\% \text{ d}^{-1} \text{ stage}^{-1}$) (D) Heinle graph for DECSURV (copepodite survivorship decreased $2\% \text{ d}^{-1} \text{ stage}^{-1}$)

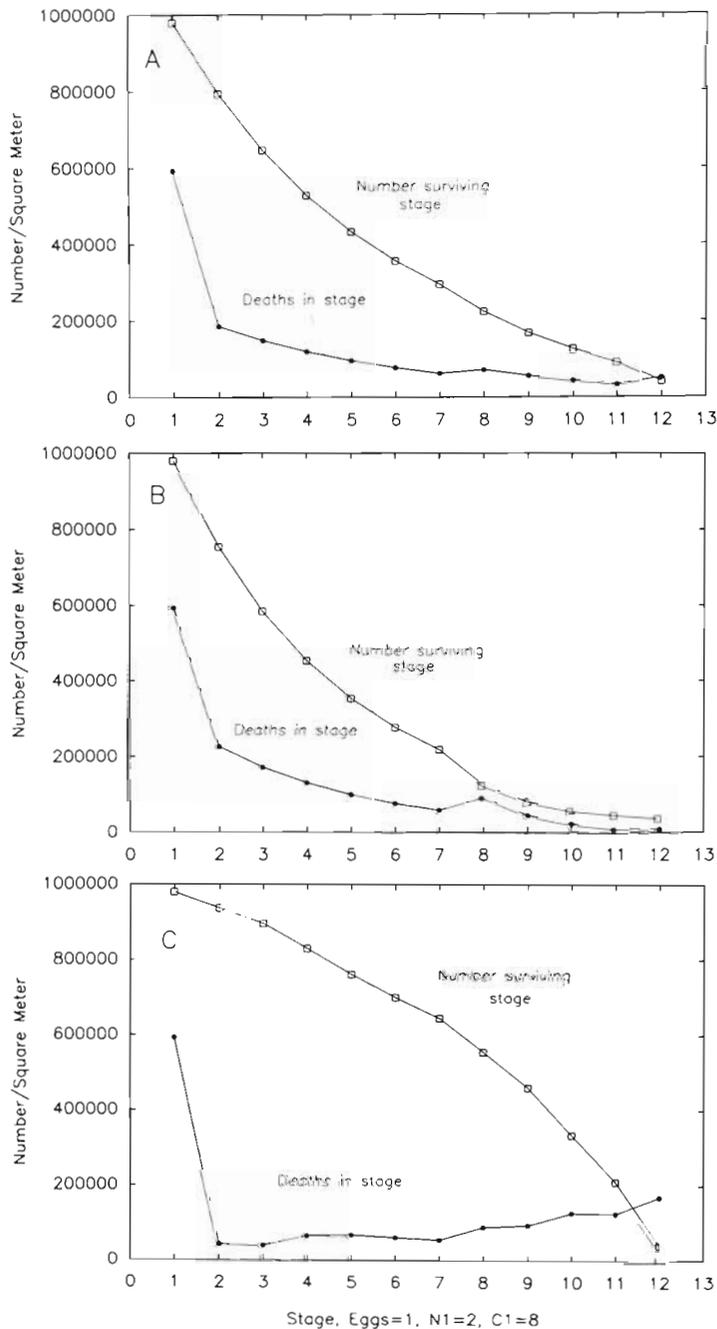


Fig. 13. Survivorship curves for (A) CONSTSURV, (B) INCSURV, and (C) DECSURV

copepodite stages. Both survivorship sets were tuned to produce the same final survivorship as MALANGEN. Heinle graphs (Fig. 12C, D) underestimated C4 duration by 23% in INCSURV and overestimated it by 13% in DECSURV. As pointed out above, there is no *a priori* reason to expect any particular pattern of survivorship differentials among stages. Thus, without separate information on survivorship, there is no certain way to

determine the differential mortality bias of Heinle graph estimates.

We believe the examples are extreme; the bias generated in them by differential mortality among stages is important, but it does not overwhelm the value of Heinle graphs for obtaining approximate stage durations from *Calanus* populations in the field.

In addition to bias from differential mortality, stage duration estimates from Heinle graphs are affected by skewness in the spawning function. Two simulations, SKEW-EARLY and SKEW-LATE, both based on MALANGEN, demonstrate the strength of this effect. Extreme skewing of egg production to early dates in SKEW-EARLY (Fig. 14A) reduced the Heinle measure of C4 duration from 6.63 d expected from the Belehrádek function to 5.96 d (Fig. 14B), a total downward bias of 10%, or 2% more than attributable to differential mortality. In SKEW-LATE (Fig. 14C), extreme skew of spawning to late dates delayed the C3→C4 transition by 10 d relative to MALANGEN, with estimated C4 duration of 5.65 d (Fig. 14D), a total downward bias of 10.5% compared to the Belehrádek expectation of 6.31 d. The bias from skew timing of spawning is shortening of the apparent stage duration in both cases; the important feature is having one steep face on the spawning curve.

The Heinle graph procedure appears, despite the biases revealed by the model, to be useful for obtaining first order field estimates of stage duration, within about 10% apart from statistical variation in fitting the transition curves. Miller (in press) came to the same conclusion using a different model based on a set of field data for stage proportions. He did not consider bias from skewness of the spawning function, but did show that age distribution was nearly symmetrical by C1 for the univoltine population of *Neocalanus plumchrus* in the subarctic Pacific. Estimates of bias in our model examples apply to the *expected* values for the Heinle graph estimates of stage duration. Variability will be introduced to field estimates by sampling variance in proportion estimates and by spatial patchiness in the population with respect to timing of the stage progression. Miller (in press) showed for one specific case that variance of proportion estimates is a lesser source of variability than patchiness.

Simulation of 'Kimmerer experiments'

Kimmerer & McKinnon (1987) suggested a method for determination of tissue production rate in pelagic crustaceans for which a population steady state can reasonably be assumed. That assumption is clearly violated by populations of *Calanus finmarchicus*, but the method has been applied to them anyway by Peterson et al. (1991), who did exercise great caution in interpretation. The method is to collect a quantity of mixed copepodites by very gentle net tow, then to sieve off both the largest stage (held back by a coarse mesh) and the nauplii (which pass a fine mesh). Then a portion are preserved and a portion are incubated for a full day then preserved. Stage proportions are estimated by counting, and the molting rates for each stage are calculated according an equation from Peterson et al. (1991):

$$MR_j = \frac{\sum_{k=j+1}^{11} N_{k,f} - \sum_{k=j+1}^{11} N_{k,s}}{N_{j,s}}$$

where $N_{k,f}$ and $N_{k,s}$ are the estimated final and starting proportions of stages older than the stage of interest, and $N_{j,s}$ is the estimated starting proportion of that

stage. Peterson et al. (1991) also point out that if growth is continuously exponential, which it generally is to a good approximation, then growth rate can be calculated for each stage as

$$g_j = \ln \frac{W_{j+1}}{W_j} \times MR_j$$

where W_j and W_{j+1} are the weights of the successive stages.

For convenience (and credit) we term this a 'Kimmerer experiment'. The method is attractive for its simplicity. A simple modification of MALANGEN allowed simulation of Kimmerer experiments for every day of the stage progression. At the end of each day a copy was made of the A_j and D_j vectors for $j = 8, 9, 10,$ and 11 (that is, C1 to C4, equivalent to sieving out nauplii and C5), and the copies were carried through the development of the following day exactly as for the originals but with no mortality (perfect experimental technique). Then molting rates for C1 to C4 were calculated from the Peterson equation. A time series of the results from a simulated Kimmerer experiment with the MALANGEN parameter set (Fig. 15A) shows that, over the intervals in which the successive stages are abundant enough for useful counting, the expecta-

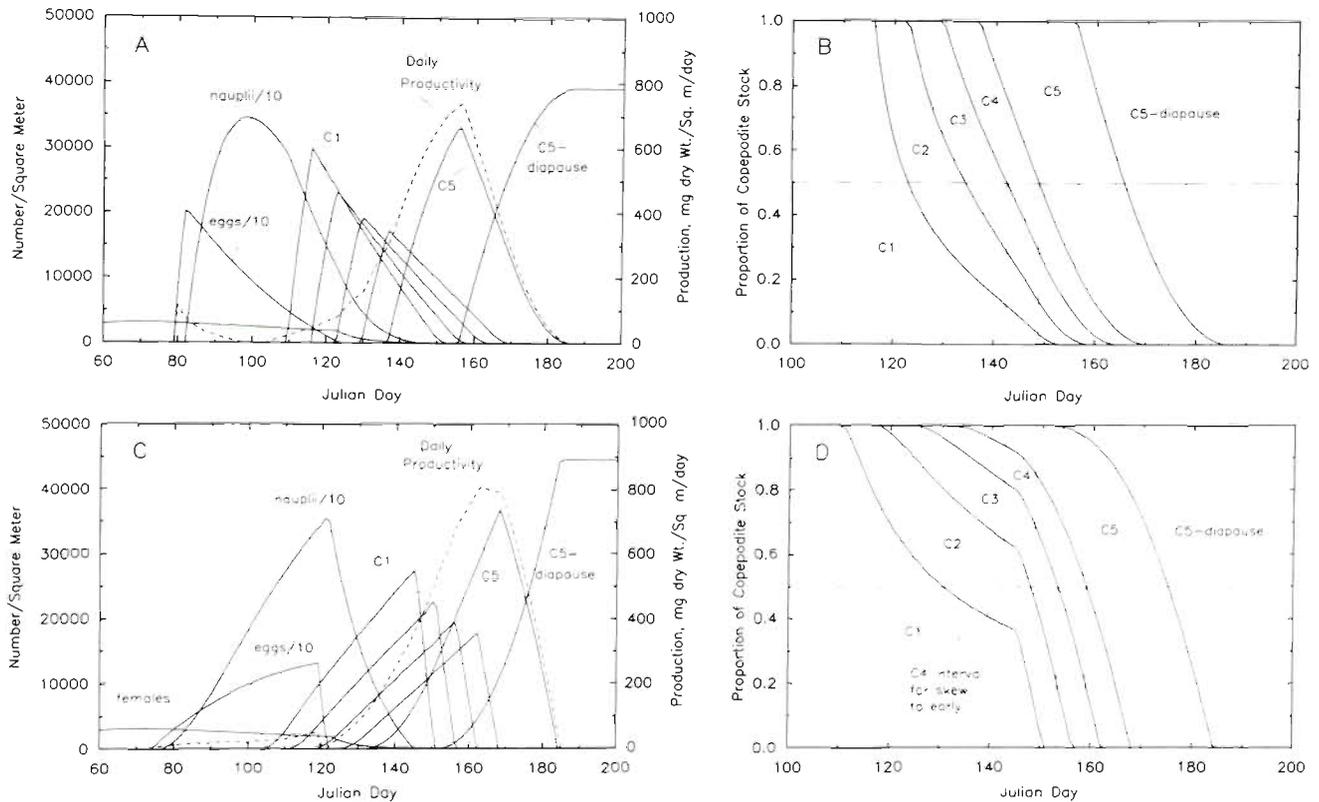


Fig. 14. (A) Output for SKREW-EARLY version of model. Skewness appears in the egg abundance curve. Total production was 11.8 g dry wt m⁻² (B) Heine graph for SKREW-EARLY. (C) Output for SKREW-LATE version of model. Skewness appears in egg abundance curve. Total production was 13.3 g dry wt m⁻². (D) Heine graph for SKREW-LATE

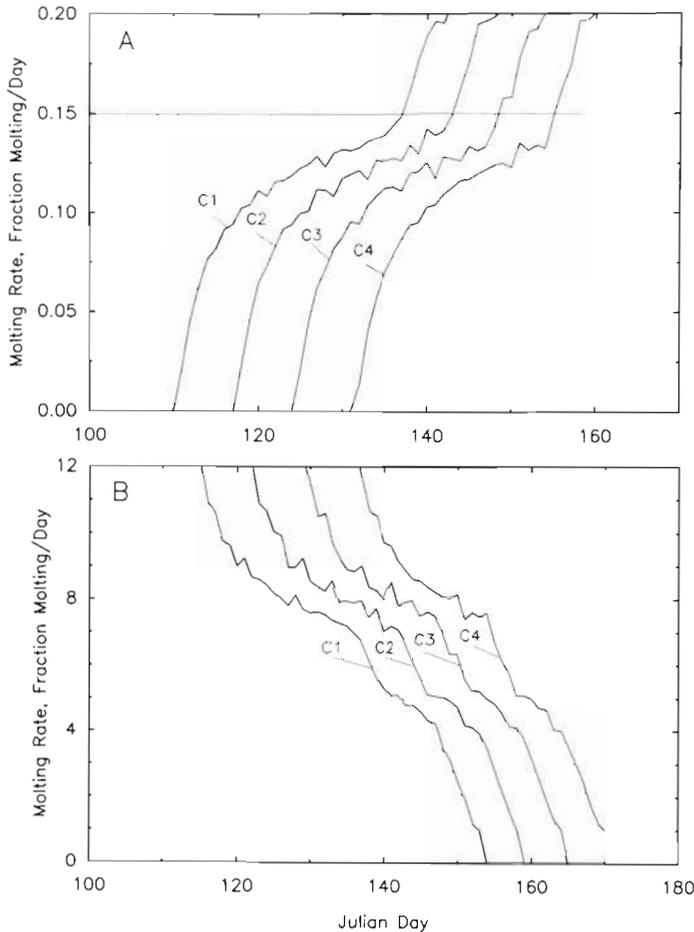


Fig. 15. (A) Molting rates, MR_j , from a 'Kimmerer experiment' performed daily using output from MALANGEN. (B) Stage duration estimates calculated as $1/MR_j$, also for MALANGEN

tions for MR_j are not very stable, rising from about 10 to about 15% over 20 d. If the MR_j are inverted to obtain estimates of stage duration (Miller et al. 1984), the expectations are also unstable (Fig. 15B), and give substantial overestimates of the Belehrádek expectations in the parts of the graph based on substantial numbers of each stage. That is because the age-within-stage distribution favors younger individuals until well past the time of the abundance peak for each stage (Fig. 8).

Next, we conducted Kimmerer experiment simulations for the SKEW-EARLY and SKEW-LATE versions of the model. In the former, MR_j estimates were biased upward by about 5% throughout the stage progression (Fig. 16A). That is because once any individuals are old enough to molt from a stage, the age-within-stage distribution favors older individuals (thanks to the prevalence of early spawning over late spawning). In the latter, MR_j were less affected (Fig. 16B), possibly because egg production in MALANGEN is already mod-

erately skewed toward late dates in the spawn (Fig. 5).

We conclude that because of changing expectations and substantial possible biases, Kimmerer experiments are unlikely to give useful information about molting rate, growth rate, or stage duration for populations with a strongly progressive cohort development. Especially considering the careful experimentation required to minimize mortality from collection and sieving, and also considering the extensive counting effort required, Kimmerer experiments are not promising for evaluation of field development rates in *Calanus* populations. The time series of MR_j estimates has a distinctive form, suggesting that a time series of such experiments might be informative despite the biases. However, the results contain no more information than a Heinle graph which can be generated by a simpler procedure.

A note on fitting of age-detailed models to data

We have worked on the problem of fitting age-within-stage models to stage abundance data using simulated data derived from MALANGEN. In doing this the egg production function was taken to be well characterized by the data used in establishing MALANGEN, and stage durations were taken to be strictly dependent upon temperature (as in the underlying model). The principal problem is fitting survivorship values, for which we developed a stage-by-stage approach. An arbitrary survivorship for N1 (S_2) was selected, daily abundances were calculated with the model until N1 appeared and again disappeared, then the fit of the model to N1 values in a simulated sample set was evaluated by summing the squares of deviations over all sample dates. Finally, S_2 was systematically varied and the model rerun for each new value until the sum of squared deviations was minimized. This was then repeated for the remaining stages. This technique reproduces the survivorships exactly when applied to data from the model at any sampling frequency.

Unfortunately, even modest sampling noise throws this simple, obvious method off course. Sampling noise was simulated by applying random Normal multipliers ranging from 0.5 to 2.0 [selected as logarithms from a logNormal distribution with $\mu = 0$, and $2\sigma = \ln(2.0)$] to each stage at each weekly sampling (Fig. 17). Deviation from the true value in fitting of S_1 is compensated

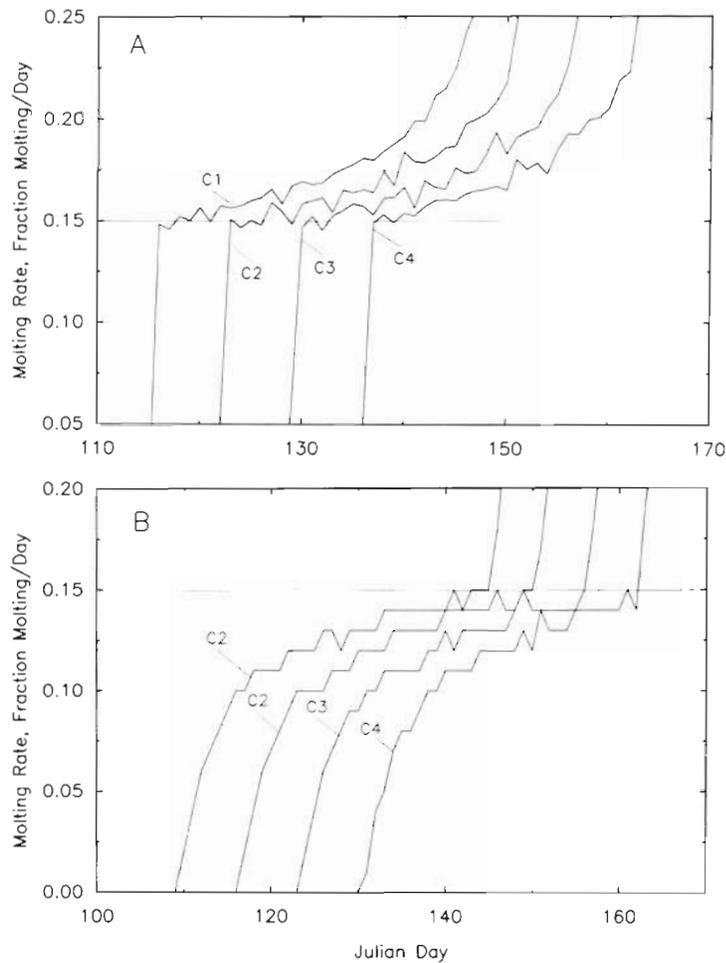


Fig. 16. Results of simulated 'Kimmerer experiment' performed on (A) SKEW-EARLY and (B) SKEW-LATE

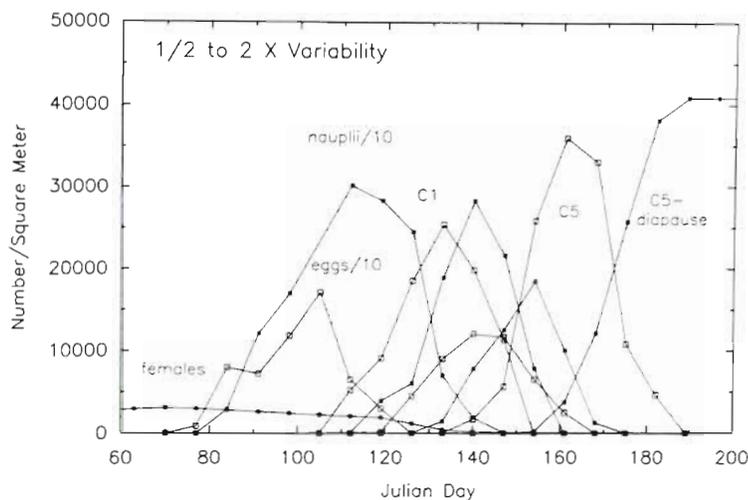


Fig. 17. Results from MALANGEN taken once each week with sampling variability applied to each stage in each sample. Variability was from random, log-normal multipliers with $2\sigma = 0.5$ to 2.0 . That is, each model value was multiplied by the antilog_e of a random Normal number with $\mu = 0$, and $2\sigma = \ln(2)$. This is a single example; every run is different

for in fitting S_2 , leading to expanding oscillation of later S_i values. By N_3 or N_4 , S_i estimates at the peaks exceed 1.0 in data with this modest noise level. This general problem has been termed 'age propagating instability' by Wood et al. (1989) and Wood & Nisbet (1991). If it is possible to estimate survivorship rates by fitting of age-within-stage models, it must be done by an overall fit of the survivorship set, S_i , to data for the whole generation. This might be accomplished by simulated annealing (e.g. Press et al. 1986) or a related Monte Carlo technique. Given the time for individual runs of the model, this is a supercomputing problem. The methods proposed by Wood & Nisbet (1991) may be applicable, but further study is required.

DISCUSSION

Selection of mortality coefficients

Estimation of secondary production by application of age-within-stage models requires adoption of stage-specific mortality coefficients. This is true of some other forms of secondary production calculation as well (Fager 1973). In testing the potential of the MALANGEN model, we have freely varied mortality to scale the stage abundance to a 'right' order of magnitude. Further thought is required as to the best rationale for doing this. Only a few studies have been conducted outlining how mortality operates in zooplankton populations (e.g. Myers & Runge 1983). The majority of information on this issue has come from fitting various models to field population data for several marine copepod species (Matthews et al. 1978, Parslow et al. 1979, Sonntag & Parslow 1981, Aksnes & Magnesen 1988). From the results, we know that mortality constants for the various life stages are site and time specific, and the estimates reflect the constraints of both the sampling designs and the mathematical methods adopted. Therefore, limited insight has been gained, as yet, from these studies for sorting out the underlying mechanisms governing mortality of copepods in natural environments.

Realizing the complexity of mortality estimation, some have suggested methods for estimation of secondary production of continuously reproducing populations (Huntley

& Boyd 1984, Kimmerer 1987, Huntley & Lopez 1992). Those are attractive, but an intrinsic interest remains in the quantification of mortality rates. Recurring quantification of mortality is needed for understanding abundance fluctuations in the annual and biennial populations of high latitude environments. Attacks on this will depend upon what are defined as important problems. For some purposes, detailed data are required. Lynch (1983) argued that a knowledge of the shape of the mortality curves as a function of age (or stage) for natural zooplankton populations would substantially improve our understanding of the evolutionary ecology of zooplankton communities. The evolution of different life history traits is likely to depend more on the age-specific pattern than on the absolute value of mortality (Charnov & Schaffer 1973). Although various survivorship schedules have been separated among insects (e.g. Birley 1977 and references therein), knowledge of these patterns for copepods is very limited. We have only begun to explore the consequences of different mortality schedules using our age-within-stage models. Data from Tande (1988) and Pedersen & Tande (1992) suggest that survivorship schedules in *Calanus* differ between constant and continuously increasing temperatures. The simulations showed that such differences in pattern are critical to annual survival and stock production. Thus, detailed studies of the response of mortality to schedules of habitat variation are in order. Exactly how to accomplish this for field populations remains to be invented.

If, on the other hand, one is primarily interested in year-to-year stock fluctuations, simpler overall estimates might suffice. We could easily simulate constant final survivorship with a huge range of stage-specific survivorship patterns (Fig. 13). Perhaps interannual variation in survivorship could be characterized from recurring studies of the abundance of resting stocks in autumn. Estimation in this case could be greatly aided by detailed acoustic or optical counting techniques, both of which are best suited to the nearly monotypic assemblages characteristic of *Calanus* resting stocks.

On fitting of the model to data

The key data for development of a model of this type are the egg production experiments from which the egg production function is described. For Malangen Fjord we have the advantage of repeated egg production determinations for *Calanus finmarchicus* throughout the adult phase of G_0 (Diel & Tande 1992). Application of age-detailed models to stocks in other places will require appropriate, local egg production studies. Data should include both daily egg production per

capita and recurring estimates of female abundance. In our model the female abundance essentially was guessed based on reasonable density and maturation timing for diapausing C5. Realism of the models can be improved with direct data on female abundance.

A more sophisticated study of the fitting of the model to simulated sampling data and to real data is in order. We only began a study of the former, and did not try to fit the model to actual data. On the whole, field data in the literature are weak on sampling of some part of the life history. Frequently, egg or naupliar data are missing altogether. In the best studies (e.g. Marshall et al. 1934, Aksnes & Magnesen 1983), nauplii and early copepodite stages are underestimated relative to later copepodite stages; the integrals of time-abundance curves for younger stages are much less than for later copepodites. The differences are several-fold, much larger than can be explained by longer durations of the later stages. For nets towed from deep to shallow, the source of this bias apparently is interaction between the deepening vertical distributions of successive stages and net clogging. The younger, shallower-living stages are undersampled relative to older ones because the filtration efficiency drops before the rising net reaches them. Marshall et al. (1934) stated the argument as follows:

‘... at this time [spring 1933] diatoms were very abundant. These would clog the meshes of the net and so reduce the amount of water filtered, particularly in the upper layers where the eggs and nauplii live.’

The same explanation must apply to the data of Grønvik & Hopkins (1984) for *Metridia longa* in Balsfjorden. They show progressively greater abundance from C1 through C3 and a further jump at C5. Again, only a modest fraction of the apparent increase is likely due to lengthening stage durations. Daytime vertical distributions in *M. longa* deepen progressively with stage of development.

In our opinion, this problem of sampling bias is likely to be the greatest obstacle to estimation of field mortality by fitting of models. There is no obvious way to compensate for the differentials in sampling bias. We suggest that thought be given to obtaining samples unbiased by clogging. Perhaps a version of Murphy & Clutter's (1972) plankton purse seine could be developed to capture early stages near the surface despite clogging. If it catches nauplii, it will also catch massive quantities of diatoms (including the glutinous *Phaeocystis pouchetti*). Thus, copepod counts will be difficult to make in the resulting samples. Another scheme might be very short tows with a multiple net, monitoring filtration on deck with flowmeters well back in the successive bags of mesh.

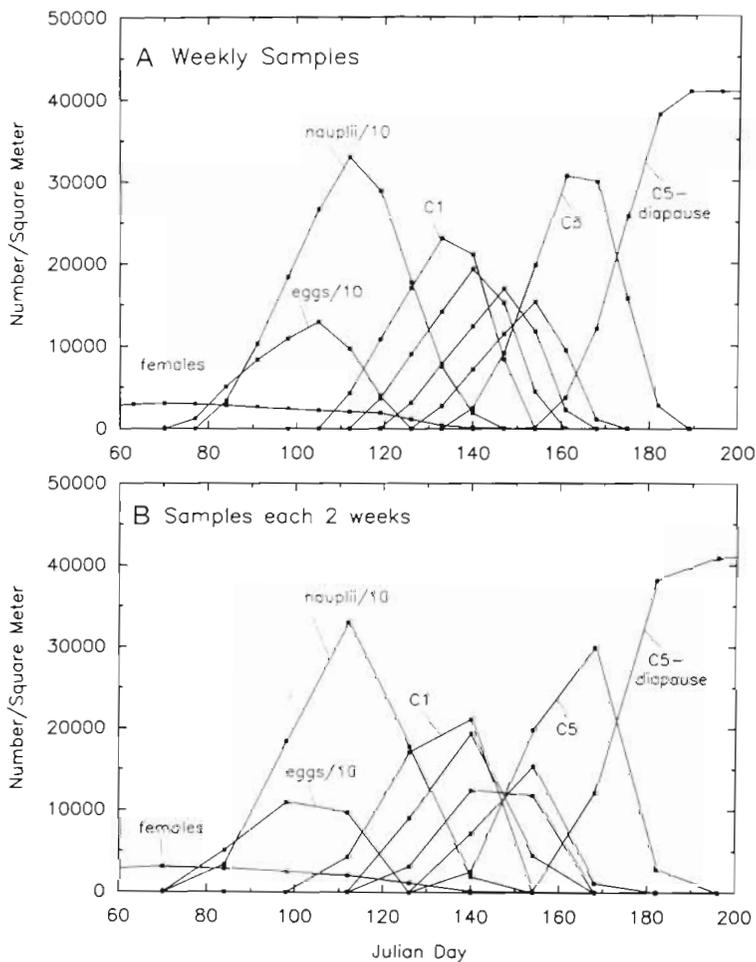


Fig. 18. Comparison of results from MALANGEN (without sampling variability) for (A) weekly and (B) biweekly sampling

Nets would be changed when filtration efficiency dropped below a criterion. This will work if clogging is not nearly instantaneous. Large capacity pumps (e.g. Taggart & Leggett 1984) are another possibility, if it can be shown that older stages are not better than younger ones at avoiding the intake. For either pumps or multiple nets, sorting of copepods from masses of diatoms will be entailed in counting the samples. This is an unpleasant but not an insuperable problem.

Fitting of age-within-stage models to real data obviously will require sampling results of high enough frequency to represent the stage progression adequately. Just from the visual impression of plots of simulations based on MALANGEN, weekly sampling (Fig. 18A) seems to be sufficient to represent most aspects of the stock progression. With sampling only every 2 wk (Fig. 18B), considerable resolution is lost. Specifically, even without field variability in sampling results, the timing of peaks is obscured, with several

stages reaching their apparent peaks on the same sampling date. This problem will be more severe for warmer habitats than for colder ones. For example, Marshall et al. (1934) found the peaks of C1, C2 and C3 in G_1 to have almost identical timing in the relatively warm Loch Striven habitat. When this occurs, it will make selection of reasonable parameters for a fit quite difficult. Extended analysis is needed of the interactive effects of sampling frequency, habitat temperature, and short-term sampling variability on parameter estimates.

Estimation of stage duration

We conclude from our simulations that stage duration can be estimated with reasonable precision by examining the temporal progression of fractional abundances of the stages, by developing data for a 'Heinle graph'. There is significant bias from confounding of the stage duration with differential mortality rates between the stages, and a small additional bias (in some cases compensating) from skewness in the timing of egg production. Our modelling of seeming worst-case differentials and skewness showed the Heinle graph estimates for duration of C4 to be biased by no more than 10%. Imprecision from sampling variability and from upstream-downstream variation in population development timing will likely be larger than this bias. At a minimum, estimates

from this approach will allow comparisons of development rates between sites and between years. An application by Durbin & Durbin (unpubl.) to the G_1 cohort of *Calanus finmarchicus* in 2 yr showed a large difference in stage duration between years. They attributed this to a difference in food availability. While food availability plays no role in our model, the model shows that the difference is not likely to arise from bias in the stage duration estimate.

Our simulations show that molting rates (and stage durations estimated from molting rates) measured by incubation experiments will not have a stable expectation for a cohort developing in the manner of the G_1 generation of *Calanus finmarchicus*. We are convinced that, for population cycles of the *Calanus* type, 'Kimmerer experiments' are not worth the investment of time and labor they require. This conclusion should not be extended to the roughly steady-state populations for which Kimmerer & McKinnon (1987) developed the technique.

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